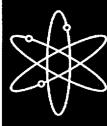


Sensitivity and Uncertainty Analyses Applied to Criticality Safety Validation



Illustrative Applications and Initial Guidance



Oak Ridge National Laboratory



U.S. Nuclear Regulatory Commission Office of Nuclear Regulatory Research Washington, DC 20555-0001



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Sensitivity and Uncertainty Analyses Applied to Criticality Safety Validation

Illustrative Applications and Initial Guidance

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ABSTRACT

This report presents the application of sensitivity and uncertainty (S/U) analysis methodologies developed in Volume 1 to the code/data validation tasks of a criticality safety computational study. Sensitivity and uncertainty analysis methods were first developed for application to fast reactor studies in the 1970s. This work has revitalized and updated the existing S/U computational capabilities such that they can be used as prototypic modules of the SCALE code system, which contains criticality analysis tools currently in use by criticality safety practitioners. After complete development, simplified tools are expected to be released for general use.

The methods for application of S/U and generalized linear-least-squares methodology (GLLSM) tools to the criticality safety validation procedures were described in Volume 1 of this report. Volume 2 of this report presents the application of these procedures to the validation of criticality safety analyses supporting uranium operations where the enrichments are greater than 5 wt %.

Specifically, the traditional $k_{\it eff}$ trending analyses are compared with newly developed $k_{\it eff}$ trending procedures, utilizing the D and c_k coefficients described in Volume 1. These newly developed procedures are applied to a family of postulated systems involving U(11)O₂ fuel, with H/X values ranging from 0 to 1000. These analyses produced a series of guidance and recommendations for the general usage of these various techniques. Recommendations for future work are also detailed.

CONTENTS

| ABS' | TRACT | iii |
|------|--|----------------------------|
| LIST | OF FIGURES | ⁄ii |
| LIST | OF TABLES | ix |
| EXE | CUTIVE SUMMARY | хi |
| ACK | NOWLEDGMENTS | ۲V |
| 1 | INTRODUCTION | |
| 2 | DEVELOPMENT OF CRITICAL EXPERIMENT DATABASE 2.1 LEU Systems 2.2 CSEWG Benchmarks 2.3 PCTR Experiments 2.4 HEUMET Systems 2.5 Rocky Flats Experiments 2.6 Russian Critical Experiments 2.7 Other Systems 2.8 Tabulation of Calculated-versus-Experimental Results and Uncertainties 1 | 5 9 11 13 |
| 3 | USE OF METHODOLOGY FOR GREATER THAN 5-WT % ENRICHED URANIUM APPLICATIONS 3.1 Description of U(11)O ₂ Systems 3.2 Sensitivity and Uncertainty Analyses of the U(11)O ₂ Systems 3.3 Traditional Trending Analysis 3.4 Trending Analysis Using D Values 3.5 Trending Analysis Using c _k Values 3.6 GLLSM Analysis of 11-wt % Systems 3.7 Summary of Trending Analyses | 19 19 25 34 38 |
| 4 | GUIDANCE FOR DETERMINATION OF APPLICABILITY AND ADVANCED TRENDING PROCEDURES 4.1 Recommendations on the Use of Trending Procedures 4.2 Trending Procedures for c _k and D Parameters 4.3 Guidance on Estimating Experimental Uncertainties 4.4 Future Developments | 47 48 48 |
| 5 | SUMMARY | 51 |
| 6 | REFERENCES | 53 |

LIST OF FIGURES

| Figu | re Page |
|------|---|
| 1 | Correlation coefficients, c_k , for U(11)O ₂ , H/X = 0 system with 102 benchmark experiments |
| 2 | Correlation coefficients, c_k , for U(11)O ₂ , H/X = 3 system with 102 benchmark experiments |
| 3 | Correlation coefficients, c_k , for U(11)O ₂ , H/X = 40 system with 102 benchmark experiments 23 |
| 4 | Correlation coefficients, c_k , for U(11)O ₂ , H/X = 500 system with 102 benchmark experiments 24 |
| 5 | D coefficients for $U(11)O_2$, $H/X = 0$ system with 102 benchmark experiments |
| 6 | D coefficients for U(11)O ₂ , H/X = 3 system with 102 benchmark experiments $\dots 27$ |
| 7 | D coefficients for U(11)O ₂ , H/X = 40 system with 102 benchmark experiments |
| 8 | D coefficients for $U(11)O_2$, $H/X = 500$ system with 102 benchmark experiments |
| 9 | Trend plot for k_{eff} versus energy of average lethargy causing fission (EALF) |
| 10 | Trend plot for k_{eff} versus energy of average lethargy causing capture (EALC) |
| 11 | Trend plot for k_{eff} versus hydrogen-to- ²³⁵ U ratio (H/X) |
| 12 | Trend plot for k_{eff} versus enrichment |
| 13 | Trend plot for k_{eff} versus D_{sum} value for the U(11)O ₂ H/X = 3 system |
| 14 | Trend plot for k_{eff} versus D_{sum} value for the U(11) O_2 H/X = 40 system |
| 15 | Trend plot for k_{eff} versus D_{sum} value for the U(11)O ₂ H/X = 500 system |
| 16 | Trend plot for k_{eff} versus c_k value for the U(11)O ₂ H/X = 0 system |
| 17 | Trend plot for k_{eff} versus c_k value for the U(11)O ₂ H/X = 3 system |
| 18 | Trend plot for k_{eff} versus c_k value for the U(11)O ₂ H/X = 40 system |
| 19 | Trend plot for k_{eff} versus c_k value for the U(11)O ₂ H/X = 500 system |
| 20 | Plot of convergence in predicted k_{eff} bias, [(e-a)/c], for U(11)O ₂ systems as a function of number of experimental groups considered |

LIST OF TABLES

| Tabl | e Page |
|------|---|
| 1 | Description of 102 benchmark experiments |
| 2 | PCTR central region experiments |
| 3 | Dimensions for Tuballoy- and nickel-reflected spheres |
| 4 | Spherical models for Rocky Flats H/U = 0.77 experiments |
| 5 | Spherical models for Rocky Flats H/U = 2.03 experiments |
| 6 | Benchmark descriptions for ORNL L-8 to L-11 experiments |
| 7 | Measured-versus-calculated k_{eff} values for 102 critical benchmarks |
| 8 | Correlation coefficients for the U(11)O ₂ systems |
| 9 | Predicted Δk bias and its standard deviation based on traditional trending procedure |
| 10 | Predicted Δk bias and its standard deviation based on D_{sum} trending procedure |
| 11 | Predicted Δk bias and its standard deviation based on c_k trending procedure |
| 12 | Predicted Δk bias and its standard deviation based on GLLSM procedure |
| 13 | Comparison of predicted Δk bias and the standard deviation of the estimated k_{eff} for various procedures . 46 |
| 14 | Comparison of predicted USL1 and USL2 values for the various procedures |

EXECUTIVE SUMMARY

Sensitivity and uncertainty (S/U) analysis and the subsequent application of generalized linear-least-squares methods (GLLSM) were developed in the 1970s primarily for application to fast reactor studies. In the 1980s these techniques were expanded in the area of reactor pressure vessel damage analyses. Recently, interest in the United States has increased in the area of cross-section evaluation, with the recognition that these methods can be used to improve the consistency of the nuclear data with respect to integral nuclear criticality experiments. Volume 1 of this report presented the methodology for application of the S/U and GLLSM techniques to the validation of data used in a criticality safety analysis.

Current generation techniques in criticality safety validation studies typically use trending analyses where the calculated effective neutron multiplication factors $(k_{\it eff})$ for sets of benchmark critical experiments are trended against physical and spectral indices, like hydrogen-to-fissile ratio (H/X), energy of average lethargy causing fission (EALF), enrichment, solution concentration, etc., to observe their patterns of variation with measured $k_{\it eff}$ values (typically 1.00). This trending allows for visual estimates of measured- versus-calculated discrepancies (i.e., biases and associated uncertainties) as a function of these spectral/physical parameters. Under this approach, the set of critical benchmarks must be shown via an expert judgement determination to be "applicable" to the particular application area intended for criticality safety studies.

Thus, a great deal of judgement is often needed in the current validation techniques in order to establish the area of applicability. Each of these traditional trending parameters like H/X and EALF can be useful in establishing possible areas of applicability; however, most systems have multiple variables and their simultaneous variation can make a definite determination of applicability difficult. The combined variations in H/X, soluble poison concentrations, reflected/unreflected geometry, enrichment, and impurity concentrations are treated poorly by using single- or even multiple-parameter trend curves. Similarly, no method is available to determine if there is sufficient coverage for the entire range of the parameter trend curves. For example, given systems with H/X values of 200, 500, 1500, is there coverage of the entire range from 200–1500? If experiments with H/X of 0 and 50 are added, is there adequate coverage from 0–1500? If two of these systems have soluble poison levels of 200 ppm and 500 ppm, is there sufficient coverage for 0–500 ppm levels of poison over the full range of H/X? These are difficult questions, with answers typically derived from expert judgement.

Conversely, the application of S/U techniques allows for a formal estimation of the applicability of each of the critical benchmarks in the set to the application area under consideration. Volume 1 of this report has presented techniques that can be utilized in a formal estimation of applicability using integral parameters based on differences in sensitivity profiles between systems (D) and the correlation coefficients between systems (c_k). The D values are defined as the following:

$$D_{n} = \sum_{i=1}^{g} | S_{nai} - S_{nei} |$$

$$D_{c} = \sum_{i=1}^{g} | S_{cai} - S_{cei} |$$

$$D_{s} = \sum_{i=1}^{g} | S_{sai} - S_{sei} |,$$

where S is the sensitivity of k_{eff} for the safety application, a, or experimental configuration, e, to $\bar{v}(n)$ for group i or to the capture (c) or scattering (s), cross sections. These parameters are useful, not only as formal estimations of critical benchmark data applicability, but additionally as trending parameters in the traditional criticality safety data validation approach.

Executive Summary

A criterion for determining the appropriate cut-off values of D and c_k for adequate "similarity" has been established. The parameters developed for use in this process are calculated by performing a sensitivity and uncertainty analysis on both the selected experiments and the applications of interest. The complete S/U analysis capability is presently limited to one-dimensional (1-D) and two-dimensional (2-D) models of the systems of interest. Future development work will hopefully eliminate this limitation.

A further development under this current work was the investigation of GLLSM. This methodology is an alternative approach to the traditional trending analysis for the determination of bias and the associated uncertainty. Physically the GLLSM is designed to minimize differences between the measured and calculated values of k_{eff} by predicting data changes based on the entire set of criticals used in the data validation process. The inputs needed for such an analysis are almost identical to the concepts presented thus far: the sensitivity coefficients, the cross-section uncertainties, and the actual calculated and measured k_{eff} values, with the addition of an estimate of the uncertainty in the measured k_{eff} values. The "data changes" that result from the application of the GLLSM can then be used to predict (via interpolation or extrapolation) the biases for *any* application determined to be similar to the benchmark area of applicability.

One of the benefits of the GLLSM approach is that, not only can the bias for a given application be estimated, the cumulative "combination" of critical benchmarks can be used to determine the convergence of the procedure. Questions that can be addressed include: *how many* experiments are needed to verify an application, and *how much* correlation to the application is necessary in order to validate the application area?

In Volume 1, the GLLSM is used to predict *how many* experiments and *how much* correlation between experiments are necessary to validate an application area. In Volume 2, the GLLSM is applied as a data validation tool for comparison with trending analyses using typical characteristic parameters as well as the D and c_k parameters.

This current report presents an illustrative application of both the S/U and GLLSM procedures to the validation of criticality safety studies for facilities processing uranium fuels with enrichments greater than 5 wt % for use in commercial power reactors. In the past, these processing facilities have been limited to enrichments at or below 5 wt %. Hence, much of the critical experiment data correspond to these lower enrichments. As a part of this study, a number of critical experiments in the 5–20-wt % range were identified as having been performed in Russia. A number of these experiments were obtained and documented as a result of this work.

As with any criticality data validation, the goal is to estimate the bias trends for ranges over which the criticality safety calculational studies are to be performed. The usefulness of S/U and GLLSM methods in validation studies is demonstrated by performing a validation of a hypothetical set of application scenarios, which consist of 14 systems, each having $U(11)O_2$ fuel with H/X values varying from 0 to 1000. The 11-wt % enrichment is chosen so the entire range of moderations, including dry, is able to achieve criticality. The data validation included traditional trending analyses, trending analyses with the D and c_k parameters developed in Volume 1, and finally the full GLLSM approach. Volume 2 explores the advantages and disadvantages of each approach and presents guidance for general use of these techniques.

The comparisons among the various trending techniques are quite interesting in that they give very different answers, depending on the particular system being analyzed. The predicted biases for various systems are in some cases up to a factor of 15 different between the various trending techniques. The primary reason for these differences is that various systems "look" very similar from the standpoint of certain parameters, but very different with respect to other parameters. In particular, the H/X and EALF parameters, frequently used in current validation efforts, predict similarity between dry systems with high enrichments (93 wt %) and dry systems with intermediate enrichments (10 wt %), while the $c_{\rm k}$ and D parameters indicate that these systems are

quite different. The net effect of trending with traditional H/X and EALF parameters is a cancellation of effects (the high-enriched systems underpredict by about 1%, the intermediate-enriched systems overpredict by about 2%), which produce a predicted bias of about 0.5% overprediction for a dry $U(11)O_2$ system. The trending with c_k and D parameters produces an estimated bias of between 1 and 2% overprediction because only the intermediate-enriched systems are predicted to be similar to the dry $U(11)O_2$ system. Although the predicted bias from these applications are all positive, the differences in magnitude are a concern, since the prudent application of trending procedures is very important in criticality validation exercises.

As a result of these observations, the following guidance was developed for the usage of the various trending techniques.

- Traditional trending procedures are acceptable for estimation of the bias and associated uncertainty if about 20 benchmarks with a c_k value of 0.8 or higher are included in the benchmark database. It is further recommended that in using the traditional k_{eff} trending procedures that systems with very different physical characteristics, such as high- versus-low enrichment, poisoned-versus-nonpoisoned, wet-versus-dry, etc., not be included together unless that specific effect is being analyzed.
- The advanced trending analyses with the D and c_k coefficients are recommended if fewer experiments than those recommended above are available. Also the primary motivation behind their use is that widely varying systems can be included in the analysis, hence making the effective extrapolation of the existing data more meaningful. The procedure will allow for automatic collecting of experiments that have similar D and c_k coefficients such that the predicted k_{eff} bias should be valid.
- The GLLSM procedure is recommended if existing experiments with c_k values of 0.8 or higher with respect to the area of application are not available. These experiments, possibly having c_k values 0.2-0.6, are not similar to the application; but they are still relevant. A suite of such experiments can be expected to perform quite well using a GLLSM technique if they are suitably chosen to highlight different aspects of the given application system. Under these conditions, the use of GLLSM for validation is possible, although it may still be desirable to initiate an experimental program to obtain additional data. This procedure also allows for the inclusion of noncritical experiments and non- k_{eff} responses in the validation exercises. Further development is required before these GLLSM techniques can become fully functional in a routine criticality safety validation study.

Additional work is needed to enable full use of these methodologies by the criticality safety practitioner. Of immediate importance is the need to extend the sensitivity analysis methodologies to the realm of three-dimensional Monte Carlo codes typically used in criticality safety analysis. Improved sensitivity information can also be obtained by including the effects of the problem-specific cross-section processing. Additional applications need to be explored to gain further confidence in the guidance developed herein and aid in the development of standard processes that can be more easily implemented and applied by criticality safety analysts. Methods to enable full implementation of GLLSM need particular attention. Finally, cross-section uncertainty information is currently rather limited and the need exists to obtain uncertainty data for a number of nuclides important to criticality safety.

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1 INTRODUCTION

Sensitivity and uncertainty (S/U) analysis and the subsequent application of generalized linear-least-squares methods (GLLSM) were developed in the 1970s, primarily for application to fast reactor studies. ^{1,2} In the 1980s they were expanded in the area of reactor pressure vessel damage analyses. ³ Recently interest in the United States has increased in the area of cross-section evaluation with the recognition that these methods can be used to improve the consistency of the nuclear data with respect to some non-criticality integral experiments. ⁴

Current-generation techniques in criticality safety validation studies⁵ use trending analyses where the calculated effective neutron multiplication factors (k_{eff}) values for sets of benchmark critical experiments are trended against physical and spectral indices, like hydrogen-to-fissile ratio (H/X), energy of average lethargy causing fission (EALF), enrichment, or solution concentration. This trending allows for visual estimates of measured-versus-calculated discrepancies (i.e., biases) as a function of these spectral/physical parameters. Under this approach, a set of critical benchmarks must be shown via an expert judgement determination to be "applicable" to the particular application area intended for criticality safety studies.

Conversely, the application of S/U techniques allows for a formal determination of the applicability of each of the critical benchmarks in the set to the application area under consideration. Volume 1 of this report has presented techniques that can be utilized in a formal determination of applicability using integral parameters based on differences in sensitivity profiles between systems (D) and the correlation coefficients between systems (c_k) . These D values are defined as the following:

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where S is the sensitivity of $k_{\it eff}$ for the safety application, a, or experimental configuration, e, to $\bar{\nu}(n)$ for group i or to the capture (c) or scattering (s) cross sections. These parameters are useful, not only as formal determinations of critical benchmark data applicability, but additionally as trending parameters in the traditional criticality safety data validation approach.

The results from such a S/U analysis can also be directly incorporated into a full GLLSM procedure quite readily, in that only additional information on the measurement uncertainties need be added to perform the analyses. The GLLSM approach described in Volume 1 of this report⁶ was used as a research tool to understand the systematics of the D and c_k parameters previously mentioned. These GLLSM results of Volume 1 supported definition of general ranges in the D and c_k parameters which indicate applicability. Based on these studies, systems with c_k values of 0.80 and higher, the sum of the D values less than 1.2, and individual D values of 0.40 and lower, can be used to designate similar systems. The use of the GLLSM procedures in a criticality safety data validation exercise has the advantage of reducing the underlying uncertainties in the criticality safety limits, which can allow for enhanced operational flexibility with little or no safety implications. Additionally, application of the GLLSM procedure should allow for the inclusion of non- k_{eff} responses (e.g., reaction rates) and even noncritical (e.g., subcritical and dosimetry) experimental data into the data validation procedure.

Introduction Section 1

This current report presents an application of both the S/U and GLLSM procedures to the validation of criticality safety studies for facilities processing uranium fuels with enrichments greater than 5 wt % for use in commercial power reactors. In the past, these processing facilities have been limited to enrichments at or below 5 wt %. Hence, much of the critical experiment data correspond to these lower enrichments. As a part of this study, a number of critical experiments in the 5–20-wt % range were identified as having been performed in Russia. Several of these experiments were obtained and documented as a result of this work. The advantage of having these data available is so that the effect of inclusion or exclusion of these data from the benchmark database can be studied.

As with any criticality data validation, the goal is to estimate the bias and associated uncertainty for ranges over which the criticality safety calculational studies are to be performed. The use of S/U and GLLSM methods in these validations studies will be demonstrated herein by performing a validation for a hypothetical set of application scenarios, which consist of 14 systems, each having $U(11)O_2$ fuel with H/X values varying from 0 to 1000. The 11-wt % enrichment is chosen so the entire range of moderations, including dry, is able to achieve criticality. The data validation will be performed using both the traditional trending analyses, trending analysis with the D and c_k parameters, and finally the full GLLSM approach. Advantages and disadvantages of each approach will be explored. Guidance for general use of these techniques will then be given.

The database of 102 critical benchmark experiments used in this study will be described in Section 2. The application of the S/U and GLLSM techniques to the data validation of the hypothetical 11-wt % systems will be discussed in Section 3. Finally, guidance on the general application of these techniques to other situations will be presented in Section 4.

1.1 BACKGROUND

This volume applies a series of newly developed data validation techniques to criticality safety applications. Before launching into the specific analyses, a summary of a current approach is needed. Certainly there are many possible approaches, but the approach briefly described herein is representative of those currently in use at Oak Ridge National Laboratory.⁵

The data validation exercises are most appropriately named trending analyses, since the calculated values of $k_{\rm eff}$ for a set of benchmark critical experiments are plotted against a single parameter, x, typically H/X, EALF, leakage, concentration, etc., to determine a linear fit of $k_{\rm eff}$ with respect to that parameter. The entire set of calculated $k_{\rm eff}$ values is then averaged, and a "pooled" standard deviation, $\sigma_{\rm p}$, (the standard deviation of the entire data set about the average) determined. A bias, $\beta(x)$, is then determined as the difference between the linear fit and the measured $k_{\rm eff}$ value. The bias is a function of the x parameter since the linear fit to $k_{\rm eff}$ can and typically does have a non-zero slope. An uncertainty in this predicted bias is assumed to be the uncertainty in the pooled data times a statistical multiplier, F. In terms of these quantities two upper subcritical limit (USL) parameters are defined:

USL1(x) = 1 -
$$\Delta k_m$$
 - $F_1 \sigma_p$ + $\beta(x)$, or alternately
USL2(x) = 1 - $F_2 \sigma_p$ + $\beta(x)$,

where Δk_m is an additional margin to ensure subcriticality, and F_2 is a more stringent multiplier than F_1 . If the analysis finds $\beta(x)$ to be positive at any point in the trending, the methodology of Ref. 5 sets $\beta(x) = 0$ such that no credit is provided for a positive bias (over prediction of k_{eff}).

Section 1 Introduction

This work has applied these currently used validation methods, but replaced the standard parameters with the D and c_k values developed in Volume 1 of this report. These techniques combine the validation and determination of areas of applicability tasks, since the D and c_k parameters are utilized for both operations. Similarly, the results from an application of the GLLSM tools to a data validation problem is simply a prediction of the bias and its uncertainty, which can also be utilized in the safety analysis procedures described above.

2 DEVELOPMENT OF CRITICAL EXPERIMENT DATABASE

The key to any criticality safety data validation procedure is the generation of a database of critical benchmark experiments that typically covers a broad range of systems that are in some way similar to the particular application(s) of interest. Several studies^{5,7-10} have reported the generation of large databases of critical experiments and corresponding code input. The inputs for the benchmark experiments are typically generated corresponding to a three-dimensional (3-D) geometry model. Currently, only 1-D and 2-D sensitivity tools are available, which necessitates the conversion of many of these inputs to 1-D or 2-D models. The impact of the revised models on the sensitivities is expected to be small, based on selected studies of this effect.

This section will briefly describe each of the 102 critical benchmark systems included in this analysis. In each case a summary of the problem materials and dimensions will be given, along with a description of the calculation procedures followed. These problems can be broken into several groups, which include: 14 low-enriched uranium (LEU) oxide or fluoride systems with 2–5-wt % uranium fuel and paraffin or stereotex moderators, 14 Cross-Section Evaluation Working Group (CSEWG) data testing benchmarks (a full range of enrichments from about 1 to 93 wt %, dry to fully moderated), the 11 Physical Constants Testing Reactor (PCTR) k_ experiments on 2-wt % uranium fuel, 9 high-enriched uranium metal experiments (HEUMET), 8 low-H/X Rocky Flats experiments (RF low H/X) with uranium enrichments of 4.5 wt %, 36 Russian experiments, which consist of various configurations of solution tanks and fuel rod arrays with ²³⁵U enrichments ranging from 5 to 89 wt %, and 10 various experiments that include HISS, UH3, and LXX configurations. The complete list of experiments, their enrichments, H/X values, and a brief description are given in Table 1. The following sections give more detailed information on each system.

2.1 LEU Systems

This group of experiments consists of 2–5-wt % uranium oxide and fluoride systems. The UF₄-paraffin systems (experiments Nos. 1–8 in Table 1) are 2- and 3-wt % uranium systems. The critical dimensions for the unreflected homogeneous spherical systems are radii of 44.91, 38.50, 36.38 36.36, 37.67, 49.65, 36.64 and 28.78 cm, respectively. Even though these systems were cuboid systems, the critical spherical dimensions were obtained from Ref. 11 and were converted from measured critical bucklings and extrapolation distances. A total of six experiments with $U(5)_3O_8$ and stereotex were also analyzed (experiment Nos. 9–14 in Table 1). These six criticals were modeled as homogeneous spheres with radii of 35.57, 33.82, 32.17, 32.94, 33.55, and 41.05 cm, respectively. These radii differ slightly from the estimates reported in Ref. 12; however, they were obtained by buckling conversions based on the reported actual critical dimensions and extrapolation distances.

2.2 CSEWG Benchmarks

This group of experiments (15–19, 30–34, 46–48, and 51–52 of Table 1) consists of a number of benchmarks developed over a number of years by CSEWG. These benchmarks are documented in the CSEWG benchmark book; however, they will be described briefly here. GODIVA is a 93-wt %-enriched bare metal sphere with radius of 8.741 cm, and a measured eigenvalue of 1.00 ± 0.003 . Big-10 is a cylindrical system consisting of a uranium-metal core, averaging 10 wt $\%^{235}$ U, reflected by depleted-uranium metal. The core has an homogeneous axial region surrounded by interleaved plates of high-enriched uranium and natural uranium such that the average 235 U content is uniform. An equivalent spherical system was modeled with a core radius of 30.48 cm and a reflector radius of 45.72 cm. The uncertainty in the measured value of k_{eff} is approximately ± 0.003 . The ZPR 3.11 assembly is a U(14) metal and steel assembly with a depleted uranium (DU) metal and steel reflector.

Table 1 Description of 102 benchmark experiments

| | Enrichment | | | |
|----------|------------|------|---|-----|
| Exp. No. | (wt %) | H/X | Application description | Ref |
| 1 | 2 | 195 | u(2)f4-ch2 h/x = 195 | 11 |
| 2 | 2 | 294 | u(2)f4 - ch2 h/x = 294 | 11 |
| 3 | 2 | 406 | u(2)f4-ch2 h/x = 406 | 11 |
| 4 | 2 | 496 | u(2)f4-ch2 h/x = 496 | 11 |
| 5 | 2 | 614 | u(2)f4-ch2 h/x = 614 | 11 |
| 6 | 2 | 972 | u(2)f4-ch2 h/x = 972 | 11 |
| 7 | 3 | 133 | u(3)f4-ch2 h/x = 133 | 11 |
| 8 | 3 | 277 | u(3)f4-ch2 h/x = 277 | 11 |
| 9 | 5 | 147 | u(5)308-h20 h/x = 147 | 12 |
| 10 | 5 | 245 | u(5)308-h20 h/x = 245 | 12 |
| 11 | 5 | 320 | u(5)308-h20 h/x = 320 | 12 |
| 12 | 5 | 396 | u(5)308-h20 h/x = 396 | 12 |
| 13 | 5 | 503 | u(5)308-h20 h/x = 503 | 12 |
| 14 | 5 | 757 | u(5)308-h20 h/x = 757 | 12 |
| 15 | 93 | 0 | godiva-u(93) | 13 |
| 16 | 1.3 | 306 | bapl-1 u(1.3) lattice $h/x = 306$ | 13 |
| 17 | 1.3 | 382 | bapl-2 u(1.3) lattice $h/x = 382$ | 13 |
| 18 | 1.3 | 515 | bapl-3 u(1.3) lattice $h/x = 515$ | 13 |
| 19 | 10 | 0 | big-10 u(n)-refl | 13 |
| 20 | 93 | 0 | hiss(hug) u(93)-c-b $c/x = 303$ | 19 |
| 21 | 98 | 0 | u(98) h2o-refl | 15 |
| 22 | 93 | 0 | u(93) 2-u(n)-refl | 15 |
| 23 | 93 | 0 | u(93) 3-u(n)-refl | 15 |
| 24 | 93 | 0 | u(93) 4-u(n)-refl | 15 |
| 25 | 93 | 0 | u(93) 5-u(n)-refl | 15 |
| 26 | 93 | 0 | u(93) 7-u(n)-refl | 15 |
| 27 | 93 | 0 | u(93) 8-u(n)-refl | 15 |
| 28 | 93 | 0 | u(93) 11-u(n)-refl | 15 |
| 29 | 93 | 0 | u(93) 8-ni-refl | 15 |
| 30 | 93 | 1378 | orn1-1 u(93)o2(no3)2-h2o h/x = 1378 | 13 |
| 31 | 93 | 1177 | orn1-2 $u(93)o2(no3)2-h2o-b h/x = 1177$ | 13 |
| 32 | 93 | 1033 | ornl-3 u(93)o2(no3)2-h2o-b h/x = 1033 | 13 |
| 33 | 93 | 972 | orn1-4 u(93)o2(no3)2-h2o-b h/x = 972 | 13 |

Table 1 (continued)

| | <u></u> | Table I (co. | ininoon) | |
|----------|------------|--------------|---|------|
| | Enrichment | | | |
| Exp. No. | (wt %) | H/X | Application description | Ref. |
| 34 | 93 | 1835 | ornl-10 u(93)o2(no3)2-h2o h/x = 1835 | 13 |
| 35 | 1.157 | 322 | pctr $1.157\% \ 3.73 \ h/x = 322$ | 14 |
| 36 | 1.071 | 353 | pctr $1.071\% \ 3.78 \ h/x = 353$ | 14 |
| 37 | 1.006 | 381 | pctr 1.006% $3.83 \text{ h/x} = 381$ | 14 |
| 38 | 1.071 | 545 | pctr $1.071\% 5.84 \text{ h/x} = 545$ | 14 |
| 39 | 1.157 | 518 | pctr $1.157\% 5.99 \text{ h/x} = 518$ | 14 |
| 40 | 1.006 | 619 | pctr $1.006\% 6.23 \text{ h/x} = 619$ | 14 |
| 41 | 1.157 | 619 | pctr $1.157\% 6.90 \text{ h/x} = 619$ | 14 |
| 42 | 1.006 | 596 | pctr 1.006% $6.95 \text{ h/x} = 596$ | 14 |
| 43 | 1.071 | 667 | pctr 1.071% $7.14 \text{ h/x} = 667$ | 14 |
| 44 | 1.006 | 748 | pctr 1.006% $7.52 \text{ h/x} = 748$ | 14 |
| 45 | 1.157 | 650 | pctr $1.157\% 7.52 \text{ h/x} = 650$ | 14 |
| 46 | 14 | 0 | zpr-3/11 u(14)-fe-ni-cr-mn u(d)-refl | 13 |
| 47 | 21 | 0 | zpr-3/12 u(21)-c u(d)-refl | 13 |
| 48 | 16 | 0 | zpr-6/6a u(16)-na-fe u(d)-refl | 13 |
| 49 | 93 | 0 | uh3 ni-refl | 18 |
| 50 | 93 | 0 | uh3 u(n)-refl | 18 |
| 51 | 1.3 | 251 | trx-1 u(1.3) lattice h/x = 251 | 13 |
| 52 | 1.3 | 429 | trx-2 u(1.3) lattice $h/x = 429$ | 13 |
| 53 | 93 | 76 | 17 u(93)o2f2-h2o h2o-refl h/x = 76 | 20 |
| 54 | 93 | 1110 | 18 u(93)o2f2-h2o h/x = 1110 | 20 |
| 55 | 93 | 1390 | 19 u(93)o2f2-h2o h/x = 1390 | 20 |
| 56 | 93 | 126 | 110 u(93)o2f2-h2o h2o-refl h/x = 126 | 20 |
| 57 | 93 | 1270 | 111 u(93)o2f2-h2o h2o-refl h/x = 1270 | 20 |
| 58 | 5 | 405 | sheba 2 u(5)o2f2-h2o h/x = 405 | 15 |
| 59 | 4.5 | 17 | rfc 4.5% h/u=.77 met h/x = 17 | 16 |
| 60 | 4.5 | 17 | rfc 4.5% h/u=.77 met h/x = 17 | 16 |
| 61 | 4.5 | 17 | rfc 4.5% h/u=.77 met h/x = 17 | 16 |
| 62 | 4.5 | 17 | rfc 4.5% h/u=.77 met h/x = 17 | 16 |
| 63 | 4.5 | 70.7 | rf 4.5% h/u= 2.03 h/x = 70.7 | 17 |
| 64 | 4.5 | 120.1 | rf 4.5% h/u= 2.03 h/x = 120.1 | 17 |
| 65 | 4.5 | 45 | rf 4.5% h/u=2.03 h/x = 45 | 17 |
| 66 | 4.5 | 45 | rf 4.5% h/u= 2.03 h/x = 45 | 17 |
| 67 | 89 | 92 | U(89)o2(no3)2 abrod $p = 0$ | 7, 8 |
| 68 | 89 | 92 | U(89)o2(no3)2 abrod $p = 6$ | 7, 8 |
| 69 | 89 | 92 | U(89)o2(no3)2 abrod p = 7 | 7, 8 |

Table 1 (continued)

| | Enrichment | | | |
|----------|------------|-----|-------------------------------------|------|
| Exp. No. | (wt %) | H/X | Application description | Ref. |
| 70 | 89 | 92 | U(89)o2(no3)2 abrod $p = 10.57$ | 7, 8 |
| 71 | 89 | 92 | U(89)o2(no3)2 abrod $p = 12.3$ | 7, 8 |
| 72 | 89 | 92 | U(89)o2(no3)2 abrod $p = 14$ | 7, 8 |
| 73 | 89 | 92 | U(89)o2(no3)2 abrod $p = 16$ | 7, 8 |
| 74 | 89 | 375 | U(89)o2(no3)2 76 g/l 0 abrods | 7, 8 |
| 75 | 89 | 375 | U(89)o2(no3)2 76 g/l 3 abrods | 7, 8 |
| 76 | 89 | 375 | U(89)o2(no3)2 76 g/l 4 abrods | 7, 8 |
| 77 | 89 | 92 | U(89)o2(no3)2 289 g/l 0 abrods | 7, 8 |
| 78 | 89· | 92 | U(89)o2(no3)2 289 g/l 3 abrods | 7, 8 |
| 79 | 89 | 92 | U(89)o2(no3)2 289 g/l 4 abrods | 7,8 |
| 80 | 89 | 92 | U(89)o2(no3)2 289 g/l 6 abrods | 7, 8 |
| 81 | 89 | 92 | U(89)o2(no3)2 289 g/l 18 p = 4 rods | 7, 8 |
| 82 | 89 | 92 | U(89)o2(no3)2 289 g/l 36 p = 4 rods | 7, 8 |
| 83 | 89 | 92 | U(89)o2(no3)2 289 g/l 18 p = 6 rods | 7, 8 |
| 84 | 89 | 92 | U(89)o2(no3)2 289 g/l 36 p = 6 rods | 7, 8 |
| 85 | 17 | 628 | U(17)o2 34 rods 22.7 C | 7, 8 |
| 86 | 17 | 628 | U(17)o2 34 rods 218 C | 7, 8 |
| 87 | 17 | 611 | U(17)o2 74 rods 16 C | |
| . 88 | 17 | 611 | U(17)o2 74 rods 151 C | 7, 8 |
| 89 | 17 | 565 | U(17)o2 68 rods 15 C | 7, 8 |
| 90 | 17 | 565 | U(17)o2 68 rods 151 C | 7, 8 |
| 91 | 10 | 50 | U(10)02 p = 0.7 20 C | 7, 8 |
| 92 | 10 | 50 | U(10)o2 p = 0.7 166 C | 7, 8 |
| 93 | 10 | 50 | U(10)o2 p = 0.7 263 C | 7, 8 |
| 94 | 10 | 340 | U(10)o2 p = 1.4 20 C | 7, 8 |
| 95 | 10 | 340 | U(10)o2 p = 1.4 206 C | 7, 8 |
| 96 | 10 | 340 | U(10)o2 p = 1.4 274 C | 7, 8 |
| 97 | 10 | 629 | U(10)o2 p = 1.852 20C | 7, 8 |
| 98 | 10 | 629 | U(10)o2 p = 1.852 193 C | 7, 8 |
| 99 | 10 | 629 | U(10)o2 p = 1.852 263 C | 7, 8 |
| 100 | 5.64 | 973 | U(5.64)o2(no3)2 400 g/l 0 rods | 7, 8 |
| 101 | 5.64 | 973 | U(5.64)o2(no3)2 400 g/l 1 rod | 7, 8 |
| 102 | 5.64 | 973 | U(5.64)o2(no3)2 400 g/l 7 rods | 7, 8 |

The specifications in the CSEWG benchmark book gave a homogeneous spherical model with core radius of 31.61 cm and a 30-cm-thick reflector, with a reported eigenvalue of 1.0000 ± 0.0025 . ZPR 3.12 is a U(25) metal, graphite, and steel assembly with a DU metal and steel blanket. The system was modeled as specified in Ref. 13 as homogeneous concentric spheres with radii of 28.76 and 59.26 cm, respectively. The measured value of $k_{\rm eff}$ was reported as 1.0000 with no uncertainty quoted. An uncertainty in $k_{\rm eff}$ due to the use of a 1-D model was specified to be 0.0023. The final ZPR-type assembly included in this analysis was the ZPR 6.6a, which is a large 16 wt %-enriched uranium oxide, sodium and steel core with a DU and steel blanket. A 1-D homogeneous spherical model is specified that consists of core radius and blanket thicknesses of 95.67 and 33.81 cm, respectively. The measured eigenvalue was 1.0000 \pm 0.0005, with an estimated 0.05% uncertainty due to the spherical model.

The experiments labeled BAPL-1, -2 and -3 (experiments 16–18 in Table 1) are a series of experiments consisting of a lattice of water-moderated U(1.3)O₂ fuel pins in a triangular pitch. The fuel pins for each experiment are identical and have pellet, gap and clad radii of 0.4864, 0.5042, and 0.5753 cm, respectively. The respective pin pitches are 1.5578, 1.6523 and 1.8057 cm for BAPL-1, -2, and -3. These specifications were used to perform cell-weighting calculations with the resulting cell-weighted cross sections input into a 1-D cylindrical model with core and reflector dimensions of 38.1269 and 59.0 cm for BAPL-1, 36.3428 and 56.0 cm for BAPL-2, and 37.6250 and 58.0 cm for BAPL-3. The respective heights for buckling were 137.245, 129.697, and 125.535 cm. No uncertainties on the measured eigenvalues were reported.

The next family of CSEWG critical experiment benchmarks corresponds to the ORNL spheres: ORNL-1, -2, -3, -4, and -10 (experiments 30–34 of Table 1). These unreflected spheres consisted of $U(93)O_2(NO_3)_2$ and H_2O at various concentrations and with/without boron. ORNL-1, -2, and -3 have a radius of 34.595 cm, and ORNL-10 has a radius of 61.011 cm. The reported eigenvalues have an uncertainty of 0.25%.

The TRX experiments (51–52 of Table 1) are water-moderated lattices of slightly enriched (1.3%) uranium metal rods with diameters of 0.4915 cm in a triangular pitch. The pitches were set to 1.8060 and 2.1740 cm to define the TRX-1 and TRX-2 configurations, respectively. Both experiments were modeled using a cell-weighting step, which characterized the pin lattice, followed by a 1-D cylindrical model consisting of core radii of 26.2093 and 27.4419 cm, followed by a 20-cm-thick water reflector. In both cases an axial buckling height of 136.98 cm was assumed.

2.3 PCTR Experiments

This group of 11 experiments (35–45 of Table 1) were performed at the Physical Constants Test Reactor (PCTR)¹⁴ and consisted of the measurement of k_{ω} for three different enrichments near 1.0 wt %, with variable H/U values. The experiments are identified by their respective H/U values: 3.83, 6.23, 6.95, and 7.52, which have an enrichment of 1.006 wt %; 3.78, 5.84, and 7.14, which have enrichments of 1.071 wt %; and 3.73, 5.99, 6.90, and 7.52, which have enrichments of 1.157 wt %. The experimentally inferred values of k_{ω} are given in Table 2, along with their uncertainties.

Table 2 PCTR central region experiments^a

| Case No. | ²³⁵ U wt % | H/U | Exp. k |
|----------|-----------------------|------|--------|
| 1 | 1.006 | 3.83 | 0.986 |
| 2 | 1.006 | 6.23 | 0.986 |
| 3 | 1.006 | 6.95 | 0.974 |
| 4 | 1.006 | 7.52 | 0.960 |
| 5 | 1.071 | 3.78 | 1.005 |
| . 6 | 1.071 | 5.84 | 1.005 |
| 7 | 1.071 | 7.14 | 0.992 |
| 8 | 1.157 | 3.73 | 1.031 |
| 9 | 1.157 | 5.99 | 1.031 |
| 10 | 1.157 | 6.90 | 1.030 |
| 11 | 1.157 | 7.52 | 1.019 |

Experimental uncertainty is \pm 0.006 Δk (varies from 0.005 to 0.007).

2.4 HEUMET Systems

This group of nine experiments (21–29 of Table 1) include two separate measurement sets.¹⁵ The first is a single experiment consisting of a water-reflected U(98) metal sphere with radii corresponding to 6.5537 cm for the core and 33.4717 cm for the reflector. The remaining set consists of eight experiments with a U(93.5) metal (Oralloy) core, along with variable thickness reflectors of natural uranium (Tuballoy) or nickel. The dimensions are given in Table 3 for each of these situations.

Table 3 Dimensions for Tuballoy- and nickel-reflected spheres^a

| Reflector thickness (in) | Region | Outer radius (cm) | Material |
|-----------------------------|--------|----------------------|----------|
| 2.0 | 1 | 6.7820 | Oralloy |
| | 2 | 11.8620 | Tuballoy |
| 3.0 | 1 | 6.4423 | Oralloy |
| | 2 | 14.0623 | Tuballoy |
| 4.0 | 1 | 6.2851 | Oralloy |
| | . 2 | 16.4451 | Tuballoy |
| 5.0 | 1 | 6.1535 | Oralloy |
| | 2 | 18.8535 | Tuballoy |
| 7.0 | 1 | 6.0740 | Oralloy |
| | 2 | 23.8540 | Tuballoy |
| 8.0 | 1 | 6.0509 | Oralloy |
| | 2 | 26.3709 | Tuballoy |
| 11.0 | 1 | 6.0276 | Oralloy |
| | 2 | 33.9676 | Tuballoy |
| 8.0 | . 1 | 6.4627 | Oralloy |
| | 2 | 26.7827 | Ni |

[&]quot;The experimental values of k_{eff} for all cases is 1.000, with uncertainties in the experiments of ± 0.003 for reflector thicknesses of 5, 7, 8, and 11 in. and ± 0.005 for all others.

2.5 Rocky Flats Experiments

These experiments 16,17 consist of a series of measurements (59–66 in Table 1) corresponding to low-enrichment and low-moderation conditions. The fuel in all cases was U(4.5)O₂, which was packed into 1.6-mm-thick aluminum cans forming a 152-mm cube. Water was added to the oxide to achieve H/U ratios of 0.77 and 2.03. The experimental setup consisted of reflected $5 \times 5 \times 5$ arrays of these cans typically driven by a high-enriched uranium driver to achieve criticality. The drivers were a high-enriched uranium metal sphere, and high- and low-concentration uranyl nitrate solution boxes.

Eight configurations were included in this work, four each for the H/U ratios of 0.77 and 2.03. The 1-D spherical models shown in Tables 4 and 5 were generated based on conservation of the actual experimental volumes.

Table 4 Spherical models for Rocky Flats H/U = 0.77 experiments

| Table 4 Spherical models for Rocky Flats H/U = 0.77 experiments | | | | | |
|---|--------|----------------------|-------------------------|--|--|
| Configuration | Region | Outer radius (cm) | Material | | |
| RF H/U = 0.77 A | 1 | 7.334 | U metal | | |
| | 2 | 9.454 | Air | | |
| | 3 | 49.627 | Hom. $UO_2 + H_2O + Al$ | | |
| | 4 | 75.227 | Plastic reflector | | |
| RF H/U = 0.77 B | 1 | 11.958 | High conc. solution | | |
| | 2 | 14.769 | Void | | |
| | 3 | 15.015 | Steel | | |
| | 4 | 49.600 | Hom. $UO_2 + H_2O + Al$ | | |
| | 5 | 75.200 | Concrete reflector | | |
| RF H/U = 0.77 C | 1 | 12.248 | High conc. solution | | |
| | 2 | 14.769 | Void | | |
| | 3 | 15.015 | Steel | | |
| | 4 | 49.656 | Hom. $UO_2 + H_2O + Al$ | | |
| | 5 | 75.256 | Plastic reflector | | |
| RF H/U = $0.77 D$ | 1 | 13.041 | Low conc. solution | | |
| | 2 | 14.746 | Void | | |
| | 3 | 14.983 | Steel | | |
| | · 4 | 49.650 | Hom. $UO_2 + H_2O + Al$ | | |
| | 5 | 75.250 | Plastic reflector | | |

[&]quot;The experimental values are 1.000, the uncertainties in the measured values of k_{eff} are not explicitly quoted, but are estimated to be ± 0.001 .

| Configuration | Region | Outer radius (cm) | Material |
|-------------------|--------|----------------------|--|
| RF H/U = 2.03 A | 1 | 34.857 | Hom. $UO_2 + H_2O + Al + Poly (0.93 cm)$ |
| | 2 | 60.457 | Plastic reflector |
| RF H/U = 2.03 B | 1 | 32.628 | Hom. $UO_2 + H_2O + Al + Poly (2.44 cm)$ |
| | 2 | 58.228 | Plastic reflector |
| RF H/U = 2.03 C | 1 | 5.669 | U metal |
| | 2 | 9.454 | Void |
| | 3 | 45 .919 | Hom. $UO_2 + H_2O + Al$ |
| | 4 | 71.519 | Plastic reflector |
| RF H/U = $2.03 D$ | 1 | 2.340 | Void |
| | 2 | 5.670 | U metal |
| | 3 | 9.454 | Void |
| | 4 | 45.746 | Hom. $UO_2 + H_2O + Al$ |
| | 5 | 71.346 | Plastic reflector |

Table 5 Spherical models for Rocky Flats H/U = 2.03 experiments^a

2.6 Russian Critical Experiments

The Russian experiments included in this work (67–102 of Table 1) have been documented in the Benchmark Handbook⁷ and separately described and analyzed in another phase of this work.⁸ These experiments include HEU-SOL-THERM-029 (7 configurations), HEU-SOL-THERM-030 (7 configurations), HEU-SOL-THERM-031 (4 configurations), IEU-COMP-THERM-02 (6 configurations), LEU-COMP-THERM-32 (9 configurations), and LEU-SOL-THERM-005 (3 configurations). Each of these configurations was modeled in 1-D. The models and results from these 1-D studies are given in Ref. 8.

2.7 Other Systems

The first two systems (49–50 of Table 1) included in this group are the UH3-UR and UH3-NI critical assemblies. These assemblies consist of approximate spheres of enriched-uranium hydride within a 20.32-cm thick natural uranium and nickel reflectors. The average empirical formula of the hydride composition is U(93.15)H_{2.97}C_{1.11}O_{0.25}. The benchmarks are modeled as a sphere within a 20.32 cm natural uranium reflector (UH3-UR) and a 20.32 cm nickel reflector (UH3-NI). The UH3-UR core has a radius of 7.7902 cm, and the UH3-NI has a core radius of 7.7978 cm. The reported value of $k_{\rm eff}$ is 1.000 with no uncertainties quoted.

The experimental values are 1.000, the uncertainties in the measured values of k_{eff} are not explicitly quoted, but are estimated to be ± 0.001 .

The HISS benchmarks are measurements of $k_{\rm eff}$ and are based on central region experiments. The single experiment considered in this work (20 of Table 1) is the homogeneous uranium graphite (HUG) measurement. The HISS(HUG) contains HEU plus graphite with a graphite/U-235 ratio of 302, and a boron/U-235 ratio of 1.16. The measured value of $k_{\rm eff}$ for this system is 1.000 \pm 0.004.

The ORNL L-7 to L-11 experiments²⁰ are a family of criticals (53–57 of Table 1) that consist of reflected and unreflected spheres of uranyl fluoride in water. Details of the models for each of these systems are given in Table 6. The vessel wall is made of aluminum. The measured k_{eff} values for each experiment are also given in the table. No uncertainty on these values is quoted. A 20-cm-thick region of water at room temperature was assumed for the reflector.

The SHEBA experiment¹⁵ is an unreflected cylinder of $U(5)O_2F_2 + H_2O$. The sensitivity calculations for this experiment were performed with a 2-dimensional model and the SEN2 code.

| Assembly | Measured $k_{e\!f\!f}$ | Inner radius (cm) | Vessel wall thickness (cm) | H/X | Water reflector |
|----------|------------------------|----------------------|-------------------------------|-------|--------------------|
| L-7 | 1.0000 | 11.5176 | 0.16 | 76.1 | Yes |
| L-10 | 1.0000 | 11.8442 | 0.16 | 126.5 | Yes |
| L-8 | 1.0004 | 27.9132 | 0.20 | 1112 | No |
| L-11 | 0.9999 | 27.9132 | 0.20 | 1276 | Yes |
| L-9 | 1.0000 | 34.6327 | 0.32 | 1393 | No |

Table 6 Benchmark descriptions for ORNL L-8 to L-11 experiments

2.8 Tabulation of Calculated-versus-Experimental Results and Uncertainties

The calculated values of $k_{\rm eff}$ and associated sensitivity profiles for the set of 102 benchmark experiments described above were evaluated using the sensitivity analysis modules SEN1 (1-D models) and SEN2 (2-D models). The cross-section library used in the analyses was the SCALE 44-group library which is derived from a 238-group library processed from ENDF/B-V. The necessary cross-section covariances were processed from ENDF/B-V into the 44-group structure. In most cases, the value of $k_{\rm eff}$ was based on the 1-D models also described above. In certain cases, like the Rocky Flats experiments, the values of $k_{\rm eff}$ are based on 3-D Monte Carlo (KENO V.a) models, although the sensitivity results are based on 1-D models. This "substitution" is necessary because the 1-D models for these cases contain significant model bias, which should be accounted for in order for the GLLSM techniques to be valid.

In Table 7, the calculated k_{eff} values of each of the 102 benchmark experiments are presented, along with their uncertainty due to cross-section uncertainties. The uncertainties in the system specification are included in the measurement uncertainties by convention. The uncertainties in the calculated k_{eff} values were estimated using the SEN1 or SEN2 modules, which incorporates the S/U methodology presented in Vol. 1 of this report. The measured k_{eff} values are also presented, along with their estimated uncertainties. The uncertainties in the measured values are incomplete, since a great deal of work is needed in most cases to quantify these values.

Table 7 Measured-versus-calculated $k_{\rm eff}$ values for 102 critical benchmarks

| | | | | TT | |
|--------------------|-----------------|-------------------------------------|--------------------|-------------------------------------|--------------------------------|
| | Calculated | Uncertainty in calculated k_{eff} | Measured k_{eff} | Uncertainty in measured k_{eff} , | (Measured/ calculated) - 1, |
| System No. | k_{eff} value | % std. dev. | value | % std. dev. | % |
| 1 U(2)F4 - 195 | 1.0065 | 1.6 | 1.000 | 0.5 | -0.65 |
| 2 U(2)F4 - 294 | 1.0079 | 1.4 | 1.000 | 0.4 | -0.79 |
| 3 U(2)F4 - 406 | 1.0049 | 1.3 | 1.000 | 0.5 | -0.49 |
| 4 U(2)F4 - 496 | 1.0029 | 1.2 | 1.000 | 0.4 | -0.29 |
| 5 U(2)F4 - 614 | 1.0020 | 1.2 | 1.000 | 0.4 | -0.20 |
| 6 U(2)F4 - 972 | 0.9954 | 1.1 | 1.000 | 0.5 | 0.46 |
| 7 U(3)F4 - 133 | 1.0219 | 1.6 | 1.000 | 0.5 | -2.15 |
| 8 U(3)F4 - 277 | 1.0193 | 1.3 | 1.000 | 0.5 | -1.90 |
| 9 U(5)3O8 - 147 | 0.9964 | 1.3 | 1.000 | 0.4 | 0.36 |
| 10 U(5)3O8 - 245 | 0.9839 | 1.1 | 1.000 | 0.4 | 1.63 |
| 11 U(5)3O8 - 320 | 1.0077 | 1.1 | 1.000 | 0.4 | -0.76 |
| 12 U(5)3O8 - 396 | 1.0021 | 1.0 | 1.000 | 0.3 | -0.21 |
| 13 U(5)3O8 - 503 | 1.0013 | 1.0 | 1.000 | 0.3 | -0.13 |
| 14 U(5)3O8 - 757 | 1.0038 | 0.9 | 1.000 | 0.3 | -0.38 |
| 15 Godiva | 1.0041 | 1.6 | 1.000 | 0.1 | -0.41 |
| 16 Bapl-1 | 1.0017 | 1.2 | 1.000 | NR * | -0.17 |
| 17 Bapl-2 | 1.0016 | 1.1 | 1.000 | NR | -0.16 |
| 18 Bapl-3 | 1.0023 | 1.0 | 1.000 | NR | -0.23 |
| 19 Big-10 | 1.0173 | 1.7 | 1.000 | 0.3 | -1.70 |
| 20 HISS (HUG) | 1.0119 | 2.1 | 1.000 | 0.4 | -1.18 |
| 21 U(98) H2O refl. | 0.9986 | 1.4 | 1.000 | 0.5 | 0.14 |
| 22 HEUMET A | 0.9881 | 1.6 | 1.000 | 0.5 | 1.21 |
| 23 HEUMET B | 0.9865 | 1.6 | 1.000 | 0.5 | 1.37 |
| 24 HEUMET C | 0.9911 | 1.6 | 1.000 | 0.5 | 0.90 |
| 25 HEUMET D | 0.9897 | 1.6 | 1.000 | 0.3 | 1.04 |
| 26 HEUMET E | 0.9956 | 1.6 | 1.000 | 0.3 | 0.45 |
| 27 HEUMET F | 0.9966 | 1.6 | 1.000 | 0.3 | 0.34 |
| 28 HEUMET G | 0.9988 | 1.6 | 1.000 | 0.3 | 0.12 |
| 29 HEUMET H | 1.0048 | 1.8 | 1.000 | 0.5 | -0.48 |
| 30 ORNL-1 | 0.9990 | 0.8 | 1.000 | 0.3 | 0.10 |
| 31 ORNL-2 | 0.9987 | 0.8 | 1.000 | 0.3 | 0.13 |
| 32 ORNL-3 | 0.9957 | 0.8 | 1.000 | 0.3 | 0.43 |
| 33 ORNL-4 | 0.9972 | 0.8 | 1.000 | 0.3 | 0.29 |
| 34 ORNL-10 | 0.9991 | 0.9 | 1.000 | . 0.3 | 0.09 |

Table 7 (continued)

| System No. | Calculated $k_{e\!f\!f}$ value | Uncertainty in calculated k_{eff} % std. dev. | Measured $k_{e\!f\!f}$ value | Uncertainty in measured k_{eff} % std. dev. | (Measured/ calculated) - 1, % |
|---------------|--------------------------------|---|------------------------------|---|-------------------------------------|
| 35 PCTR 3.73 | 1.0342 | 1.5 | 1.031 | 0.6 | -0.31 |
| 36 PCTR 3.78 | 1.0099 | 1.5 | 1.005 | 0.6 | -0.49 |
| 37 PCTR 3.83 | 0.9897 | 1.5 | 0.986 | 0.6 | -0.37 |
| 38 PCTR 5.84 | 1.0097 | 1.3 | 1.005 | 0.6 | -0.47 |
| 39 PCTR 5.99 | 1.0370 | 1.3 | 1.031 | 0.6 | -0.58 |
| 40 PCTR 6.23 | 0.9829 | 1.3 | 0.986 | 0.6 | 0.32 |
| 41 PCTR 6.90 | 1.0293 | 1.3 | 1.030 | 0.6 | 0.07 |
| 42 PCTR 6.95 | 0.9750 | 1.3 | 0.974 | 0.6 | -0.10 |
| 43 PCTR 7.14 | 0.9971 | 1.3 | 0.992 | 0.6 | -0.51 |
| 44 PCTR 7.52 | 0.9678 | 1.2 | 0.960 | 0.6 | -0.80 |
| 45 PCTR 7.52a | 1.0225 | 1.2 | 1.019 | 0.6 | -0.34 |
| 46 ZPR 3/11 | 1.0201 | 1.7 | 1.000 | 0.3 | -1.97 |
| 47 ZPR 3/12 | 1.0138 | 1.6 | 1.000 | 0.2 | -1.36 |
| 48 ZPR 6/6a | 1.0247 | 1.9 | 1.000 | 0.1 | -2.41 |
| 49 UH3 NI | 1.0108 | 1.9 | 1.000 | NR | -1.07 |
| 50 UH3 UR | 1.0072 | 1.6 | 1.000 | NR | -0.71 |
| 51 TRX-1 | 0.9947 | 1.1 | 1.000 | NR | 0.53 |
| 52 TRX-2 | 0.9975 | 0.9 | 1.000 | NR | 0.25 |
| 53 ORNL L7 | 1.0053 | 0.9 | 1.000 | NR | -0.53 |
| 54 ORNL L8 | 1.0063 | 0.8 | 1.0004 | NR | -0.63 |
| 55 ORNL L9 | 1.0034 | 0.8 | 1.000 | NR | -0.34 |
| 56 ORNL L10 | 1.0029 | 0.8 | 1.000 | NR | -0.29 |
| 57 ORNL L11 | 1.0016 | 0.8 | 0.9999 | NR | -0.16 |
| 58 SHEBA | 1.0089 | 1.0 | 1.000 | NR | -0.88 |
| 59 RF 0.77 A | 1.0037 | 1.3 | 1.000 | 0.1 | -0.37 |
| 60 RF 0.77 B | 1.0116 | 1.1 | 1.000 | 0.1 | -1.15 |
| 61 RF 0.77 C | 1.0061 | 0.9 | 1.000 | 0.1 | -0.61 |
| 62 RF 0.77 D | 1.0107 | 1.2 | 1.000 | 0.1 | -1.06 |
| 63 RF 2.03 A | 1.0062 | 1.3 | 1.000 | 0.1 | -0.62 |
| 64 RF 2.03 B | 1.0105 | 1.0 | 1.000 | 0.1 | -1.04 |
| 65 RF 2.03 C | 1.0065 | 1.5 | 1.000 | 0.1 | -0.65 |
| 66 RF 2.03 D | 1.0001 | 1.5 | 1.000 | 0.1 | -0.01 |
| 67 HST29-1 | 1.0032 | 1.1 | 1.000 | 0.7 | -0.32 |
| 68 HST29-2 | 1.0062 | 1.1 | 1.000 | 0.6 | -0.62 |
| 69 HST29-3 | 1.0002 | 1.1 | 1.000 | 0.7 | -0.02 |
| | | | | | |

Table 7 (continued)

| System No. | Calculated k_{eff} value | Uncertainty in calculated k_{eff} , % std. dev. | Measured $k_{e\!f\!f}$ value | Uncertainty in measured k_{eff} % std. dev. | (Measured/ calculated) - 1, % | |
|-------------|----------------------------|---|------------------------------|---|-------------------------------------|--|
| 70 HST29-4 | 0.9964 | 1.1 | 1.000 | 0.7 | 0.36 | |
| 71 HST29-5 | 1.0015 | 1.1 | 1.000 | 0.7 | -0.15 | |
| 72 HST29-6 | 1.0050 | 1.1 | 1.000 | 0.7 | -0.50 | |
| 73 HST29-7 | 1.0031 | 1.1 | 1.000 | 0.6 | -0.31 | |
| 74 HST30-1 | 1.0002 | 1.0 | 1.000 | 0.4 | -0.02 | |
| 75 HST30-2 | 1.0006 | 1.0 | 1.000 | 0.3 | -0.06 | |
| 76 HST30-3 | 0.9995 | 1.0 | 1.000 | 0.3 | 0.05 | |
| 77 HST30-4 | 1.0046 | 1.1 | 1.000 | 0.6 | -0.46 | |
| 78 HST30-5 | 1.0006 | 1.1 | 1.000 | 0.6 | -0.06 | |
| 79 HST30-6 | 1.0024 | 1.1 | 1.000 | 0.6 | -0.24 | |
| 80 HST30-7 | 1.0007 | 1.3 | 1.000 | 0.6 | -0.07 | |
| 81 HST31-1 | 1.0012 | 1.1 | 1.000 | 0.5 | -0.12 | |
| 82 HST31-2 | 1.0035 | 1.1 | 1.000 | 0.6 | -0.35 | |
| 83 HST31-3 | 1.0027 | 1.1 | 1.000 | 0.6 | -0.27 | |
| 84 HST31-4 | 1.0000 | 1.1 | 1.000 | 0.7 | 0.00 | |
| 85 ICT02-1 | 0.9944 | 1.0 | 1.000 | 0.4 | 0.56 | |
| 86 ICT02-2 | 0.9927 | 1.0 | 1.000 | 0.4 | 0.74 | |
| 87 ICT02-3 | 0.9997 | 1.0 | 1.000 | 0.4 | 0.03 | |
| 88 ICT02-4 | 0.9953 | 1.0 | 1.000 | 0.4 | 0.47 | |
| 89 ICT02-5 | 0.9927 | 1.0 | 1.000 | 0.4 | 0.74 | |
| 90 ICT02-6 | 0.9915 | 1.0 | 1.000 | 0.4 | 0.86 | |
| 91 LCT32-1 | 0.9977 | 1.2 | 1.000 | 0.4 | 0.23 | |
| 92 LCT32-2 | 0.9912 | 1.3 | 1.000 | 0.4 | 0.89 | |
| 93 LCT32-3 | 0.9881 | 1.3 | 1.000 | 0.4 | 1.20 | |
| 94 LCT32-4 | 1.0095 | 1.0 | 1.000 | 0.4 | -0.94 | |
| 95 LCT32-5 | 0.9977 | 1.0 | 1.000 | 0.3 | 0.23 | |
| 96 LCT32-6 | 0.9982 | 1.0 | 1.000 | 0.3 | 0.18 | |
| 97 LCT32-7 | 1.0079 | 1.0 | 1.000 | 0.5 | -0.78 | |
| 98 LCT32-8 | 1.0020 | 1.0 | 1.000 | 0.4 | -0.20 | |
| 99 LCT32-9 | 1.0026 | 1.0 | 1.000 | 0.4 | -0.26 | |
| 100 LST05-1 | 0.9988 | 1.1 | 1.000 | 0.4 | 0.12 | |
| 101 LST05-2 | 0.9989 | 1.1 | 1.000 | 0.5 | 0.11 | |
| 102 LST05-3 | 0.9989 | 1.1 | 1.000 | 0.6 | 0.11 | |

^{*}None reported (NR)

A complete analysis of the measurement uncertainties was carried out on the first 14 experiments. The details of this analysis are described in Section 4 to serve as guidance on the generation of similar results for all systems included in a GLLSM analysis.

For the purposes of this work, a range of measurement uncertainties corresponding to 0.1 to 0.4% was assumed. Various scenarios were assumed to determine the sensitivity of the final results to the explicit values assumed for the measurement uncertainties. This procedure should be useful in establishing a standard approach toward the estimation of measurement uncertainties.

3 USE OF METHODOLOGY FOR GREATER THAN 5-WT % ENRICHED URANIUM APPLICATIONS

The GLLSM methodology, as described in Volume 1 of this document,⁶ is quite complex and may be quite daunting to the criticality safety practitioner. For this reason, this section is designed to proceed through a criticality validation exercise using standard techniques, S/U techniques, and GLLSM techniques as demonstration exercises. Each of these techniques will use the same benchmark database which consists of the 102 experiments described in the previous section. These experiments are not specifically "tailored" for the greater than 5 wt % applications; they represent a generic set which could be used for a number of different applications using these new validation techniques. The S/U and GLLSM approaches have several options, which will be examined; conclusions will be drawn from the various approaches.

3.1 Description of U(11)O₂ Systems

The illustrative application for which validation is sought is a series of $U(11)O_2$ systems with H/X values of 0, 3, 5, 10, 20, 40, 80, 200, 300, 400, 500, 600, 800, and 1000. These $U(11)O_2$ systems are "artificial" critical bare spheres generated for calculational comparison purposes. The object of this demonstration problem is the validation of the underlying cross sections and criticality code for these 14 systems.

The dimensions for the $U(11)O_2$ systems were generated via a critical radius search calculation using the SCALE/XSDRNPM module. The critical radii corresponding to the 14 unreflected systems for H/X of 0 to 1000 were 50.55, 40.00, 36.85, 31.77, 26.48, 22.34, 19.87, 18.99, 19.53, 20.39, 21.42, 22.51, 25.41, and 29.17 cm.

3.2 Sensitivity and Uncertainty Analyses of the U(11)O₂ Systems

These 14 systems were analyzed with the SEN1 sensitivity module, which generates detailed sensitivity information, as well as the uncertainty in the system k_{eff} value due to cross-section uncertainties obtained from the Evaluated Nuclear Data Files (ENDF/B). The same cross-section library (SCALE 44-group) and associated covariances used in the analysis of the critical experiments were used for the analysis of these 14 applications. Using this sensitivity and uncertainty information, the uncertainty in the calculated value of k_{eff} for each system can be generated. Also, the covariance of any two systems can easily be determined. To illustrate, Table 8 provides the uncertainty matrix for the 14 U(11)O₂ systems. This matrix gives the system uncertainty for the calculated k_{eff} on the diagonal and the correlation coefficient on the off-diagonal element. These off-diagonal elements are the c_k parameters, and can be used for defining similar systems. Such a matrix can also be generated for the entire set of benchmark experiments plus applications. In this case, an 116 × 116 matrix was generated, corresponding to the 102 criticals and 14 applications. For this study, only a 14 × 102 matrix was necessary (i.e., each row corresponding to each of the 14 applications and 102 columns corresponding to each of the benchmarks). Shown in Figures 1-4 are rows of the full matrix corresponding to the U(11)O₂ systems for H/X values of 0, 3, 40 and 500. The values plotted correspond to the c_k coefficients for all 102 systems. (Note that labels are only given to every third system. The system order is the same as shown in Tables 1 and 7.) These plots allow, at a glance, the determination of the number of systems in the benchmark dataset that are similar to the particular application. From Figure 1, it is noted that 4 systems have c_k values at or greater than 0.9 compared to the H/X = 0 system. These values indicate that, based on the conclusions from Volume 1 where five systems with ck of 0.9 or higher are suggested to ensure validation, the U(11)O2 system with an H/X value of 0 is marginally covered by this benchmark set. However, it is clear from Figure 2 that the U(11)O₂ system with an H/X of 3 is not adequately covered by this

benchmark set. The remaining two graphs, shown in Figures 3–4, indicate that the $U(11)O_2$ systems with H/X values of 40 and 500 are adequately covered by this benchmark series.

Table 8 Correlation coefficients a for the U(11)O₂ systems

| Critical system | 11%-0 | 11%-3 | 11%-5 | 11%-10 | 11%-20 | 11%-40 | 11%-80 | 11%-200 | 11%-300 | 11%-400 | 11%-500 | 11%-60 | 0 11%-80 | 00 11%-1000 |
|-----------------|--------|--------|--------|--------|--------|--------|--------|---------|-------------|---------|---------|--------|----------|-------------|
| 11% H/X = 0 | 0.0191 | | | | | | | | | | | | | |
| 11% H/X = 3 | 0.8328 | 0.0185 | | | | | | | | | | | | |
| 11% H/X = 5 | 0.7379 | 0.9818 | 0.0188 | ; | | | | | | | | | | |
| 11% H/X = 10 | 0.6011 | 0.9205 | 0.9725 | 0.0188 | ; | | | | | | | | | |
| 11% H/X = 20 | 0.4887 | 0.8409 | 0.9161 | 0.9784 | 0.0176 | i | | | | | | | | |
| 11% H/X = 40 | 0.4067 | 0.7562 | 0.8403 | 0.9253 | 0.9763 | 0.0151 | | | | | | | | |
| 11% H/X = 80 | 0.3428 | 0.6585 | 0.7392 | 0.8327 | 0.9094 | 0.9698 | 0.0128 | | | | | | | |
| 11% H/X = 200 | 0.2800 | 0.5240 | 0.5888 | 0.6760 | 0.7696 | 0.8705 | 0.9526 | 0.0106 | | | | | | |
| 11% H/X = 300 | 0.2633 | 0.4751 | 0.5315 | 0.6115 | 0.7058 | 0.8157 | 0.9148 | 0.9832 | 0.0099 | | | | | |
| 11% H/X = 400 | 0.2557 | 0.4452 | 0.4953 | 0.5687 | 0.6604 | 0.7727 | 0.8798 | 0.9668 | 0.9846 | 0.0095 | | | | |
| 11% H/X = 500 | 0.2517 | 0.4329 | 0.4688 | 0.5359 | 0.6235 | 0.7349 | 0.8453 | 0.9448 | 0.9717 | 0.9845 | 0.0091 | | | |
| 11% H/X = 600 | 0.2490 | 0.4076 | 0.4482 | 0.5097 | 0.5927 | 0.7014 | 0.8123 | 0.9200 | 0.9543 | 0.9742 | 0.9847 | 0.0089 | | |
| 11% H/X = 800 | 0.2432 | 0.3755 | 0.4076 | 0.4567 | 0.5278 | 0.6265 | 0.7331 | 0.8514 | 0.8991 | 0.9330 | 0.9576 | 0.9734 | 0.0087 | |
| 11% H/X = 1000 | 0.2353 | 0.3452 | | | | 0.5509 | | 0.7702 | | 0.8731 | 0.9097 | 0.9367 | 0.9752 | 0.0087 |

[&]quot;Note the diagonal element contains the fractional standard deviation instead of the correlation coefficient.

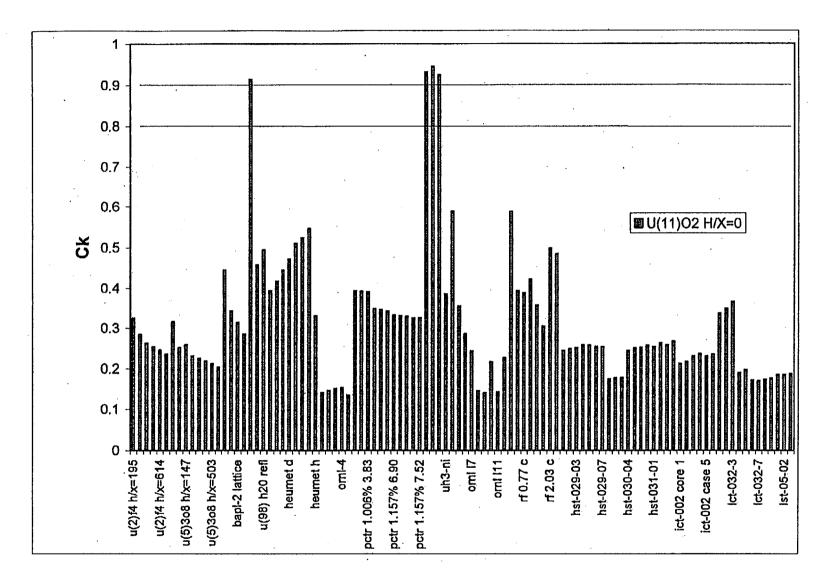


Figure 1 Correlation coefficients, c_k , for U(11)O₂, H/X = 0 system with 102 benchmark experiments

NUREG/CR-6655, Vol. 2

Application of Methodology

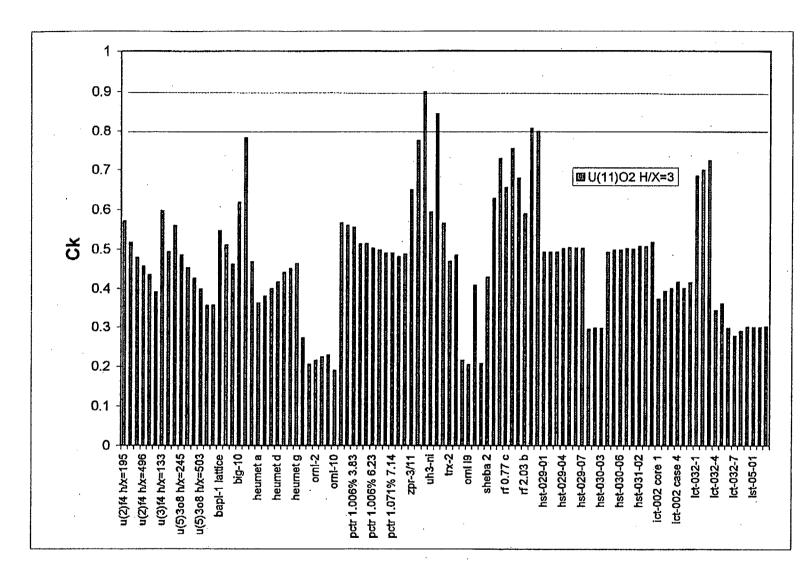


Figure 2 Correlation coefficients, c_k , for U(11)O₂, H/X = 3 system with 102 benchmark experiments

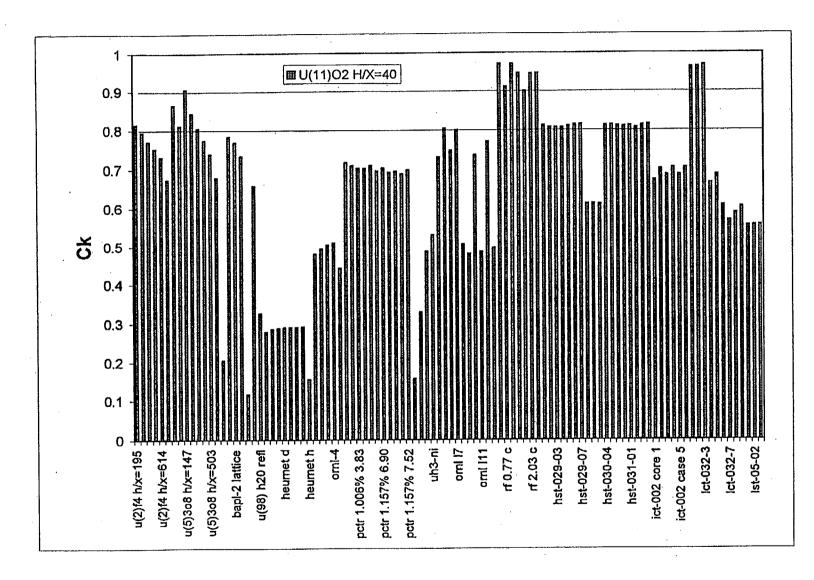
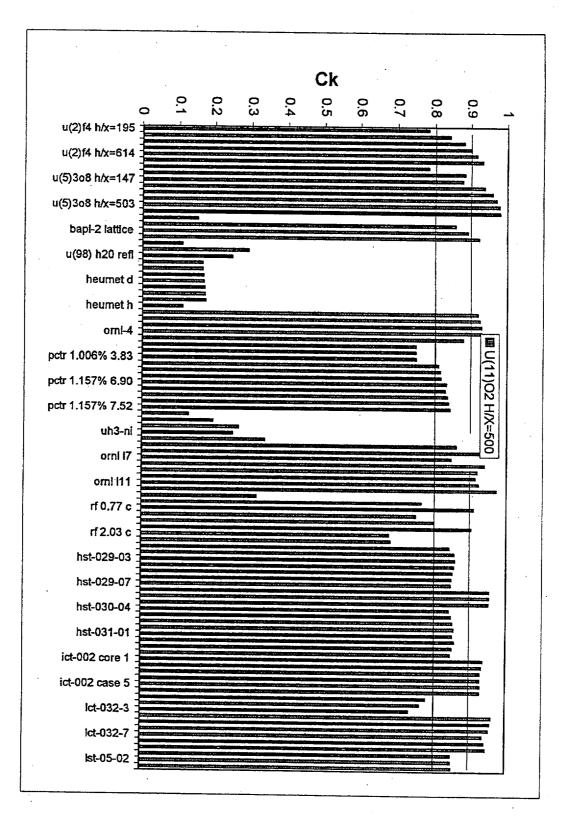


Figure 3 Correlation coefficients, c_k , for $U(11)O_2$, H/X = 40 system with 102 benchmark experiments





Similarly, a matrix of values corresponding to the D_c , D_n , and D_s values previously defined can also be generated based only on the sensitivity coefficients generated using the SEN1 module. In Figures 5–8, the sum of the D_c , D_n , and D_s values are plotted for the $U(11)O_2$ systems with H/X = 0, 3, 40, 500, respectively, to each of the 102 benchmark systems. Using this technique, the criterion for similar systems is that the sum of the D values should be less than about 1.2 (3 times the 0.4 criterion). In Figure 5, similarity is indicated for the $U(11)O_2$ H/X = 0 system and the same four systems previously determined to be similar. Figure 6 predicts similarity between the $U(11)O_2$ H/X = 3 system and only one other system. In Figures 7 and 8, the systems with H/X values 40 and 500 indicate about 10 and 40 systems, respectively, meet the less than 1.2 criterion. Hence, conclusions relative to system similarity based on these D coefficients agree with those of the preceding analysis using the c_k values.

3.3 Traditional Trending Analysis

In order to clearly show the relationship between the GLLSM techniques and the more traditional techniques for criticality safety validation, a traditional trending analysis using the same 102 benchmark experiments is presented. In Figure 9, $k_{\rm eff}$ is trended versus the energy of average lethargy causing fission (EALF). The prediction from this analysis would be a nearly constant positive bias of about 0.3%. Results shown in Figure 10 correspond to a similar parameter, energy of average lethargy causing capture (EALC). These results predict essentially the same biases as the EALF parameter and will not be further discussed.

Trend plots are also shown in Figures 11 and 12 for H/X and enrichment parameters. The H/X trend plot shows a slight trend, with the predicted Δk bias near zero for high H/X values and about +0.005 for low H/X values. The trend with enrichment is similar, but not enough data are present for the intermediate enrichments to confirm the trend.

Note that the largest variations ($\pm 2\%$) about the trend lines are seen for fast systems (i.e., the upper portion of the EALF, EALC trend plots and the lower portion of the H/X trend plot). Upon examination, it was observed for H/X = 0 systems that the predicted k_{eff} values less than unity were from the HEUMET set of criticals, while the systems with predicted eigenvalues greater than unity were from the Big-10 and ZPR sets. The combined effects of the high-versus-low enrichments and the processing of the SCALE 44-group library (collapsed using a thermal reactor spectrum) are believed to be responsible for this variation.

As a result of the trending analysis, a prediction of the Δk bias and its uncertainty can be obtained for each of the U(11)O₂ systems. Predictions using the USLSTATS procedure (Ref. 5) for systems with H/X values of 0, 3, 40, and 500 are given in Table 9.

| | .010 / 110 | motod int olds a | 110 100 01001 | 0010 00 11011011 | 00000 | additional atom | ams proc | COULT | |
|------------|------------|---------------------|---------------|---------------------|--------|---------------------|-------------------|----------------------|--|
| | H/X = | = 0 System | H/X = | = 3 System | H/X = | 40 System | H/X = 500 System | | |
| Trending | % bias | % std. dev. | % bias | % std. dev. | % bias | % std. dev. | % bias | % std. dev. of | |
| Parameters | | of biased k_{eff} | | of biased k_{eff} | | of biased k_{eff} | | biased $k_{e\!f\!f}$ | |
| EALF | 0.17 | 0.70 | 0.26 | 0.70 | 0.26 | 0.70 | 0.26 | 0.70 | |
| H/X | 0.32 | 0.71 | 0.32 | 0.71 | 0.31 | 0.71 | 0.17 | 0.71 | |

Table 9 Predicted Δk bias and its standard deviation based on traditional trending procedure

[&]quot;Standard deviations correspond to the "pooled standard deviation" as specified in Ref. 5.

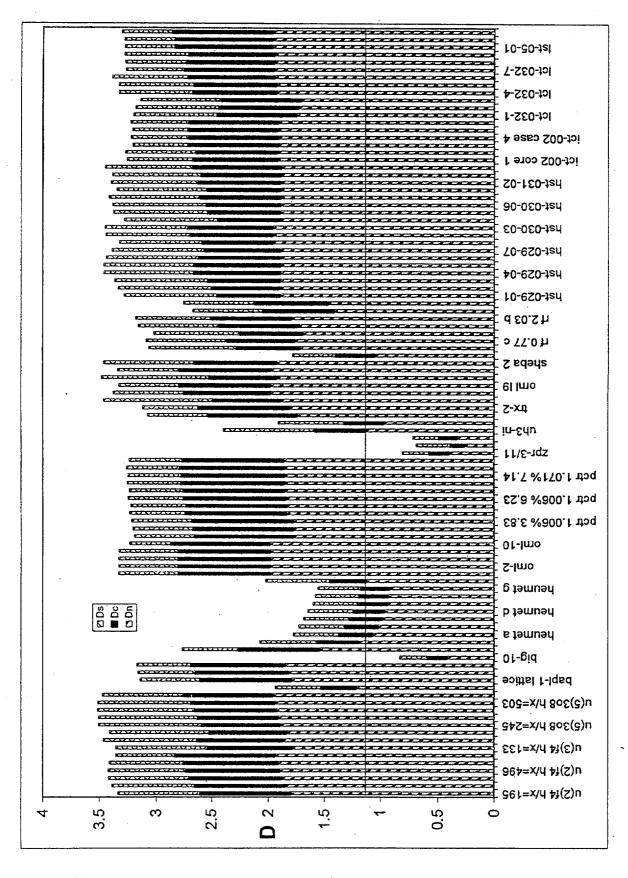


Figure 6 D coefficients for $U(11)O_2$, H/X = 3 system with 102 benchmark experiments

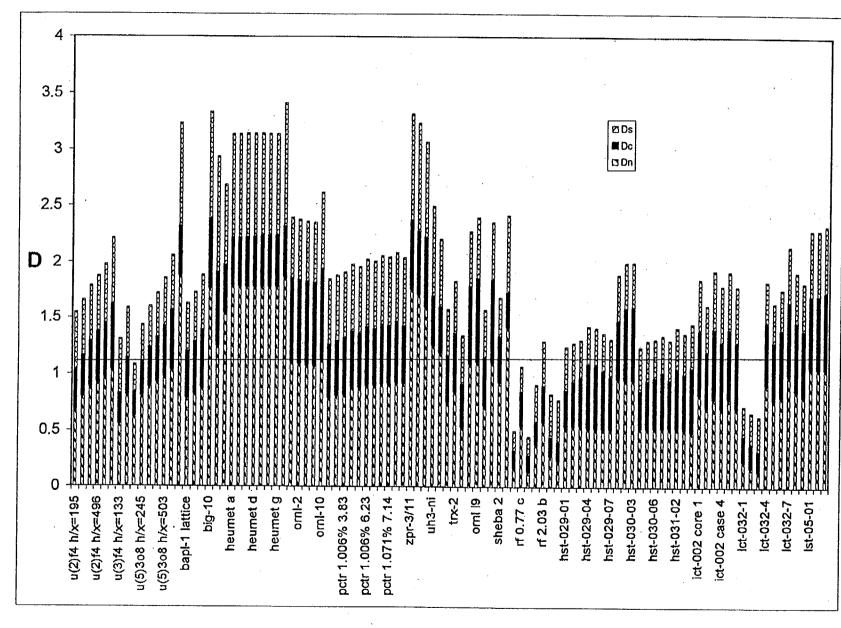
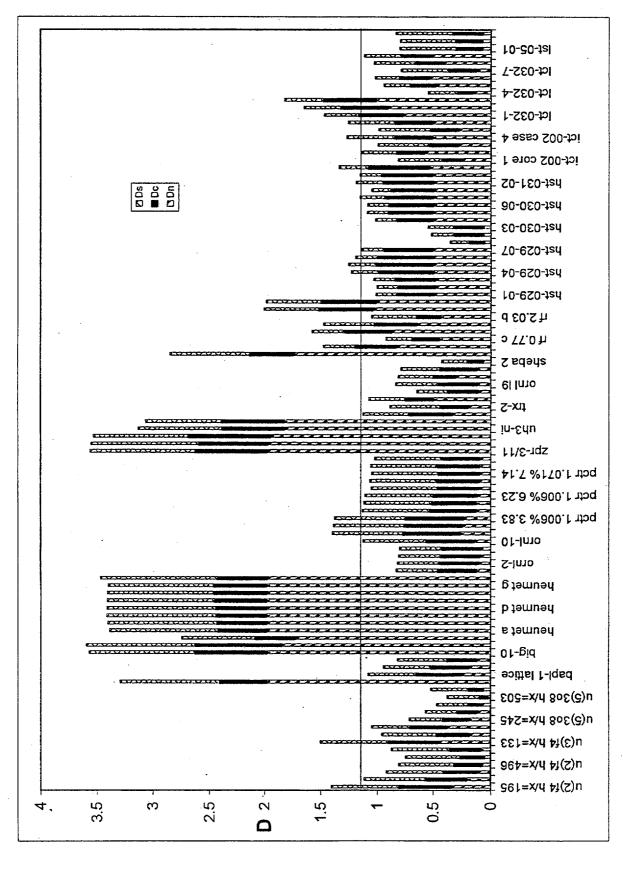


Figure 7 D coefficients for $U(11)O_2$, H/X = 40 system with 102 benchmark experiments



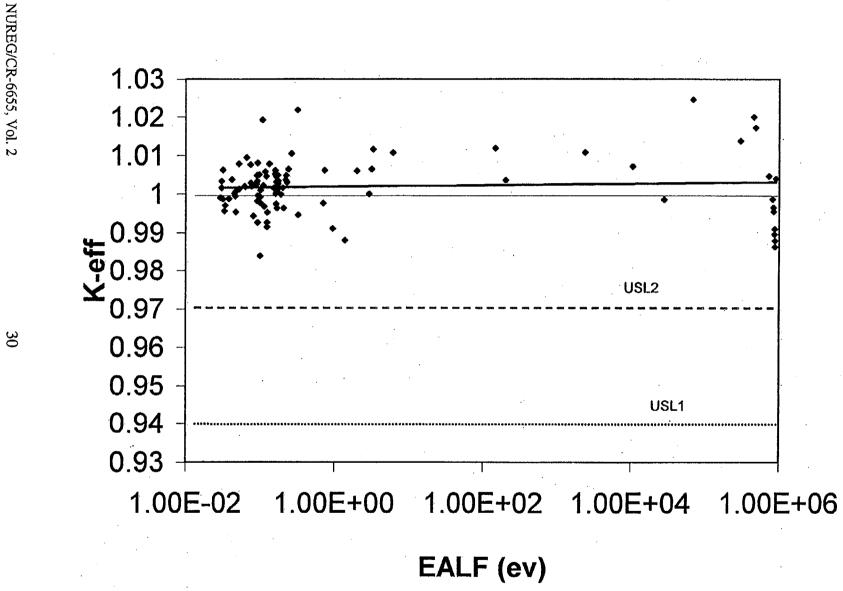
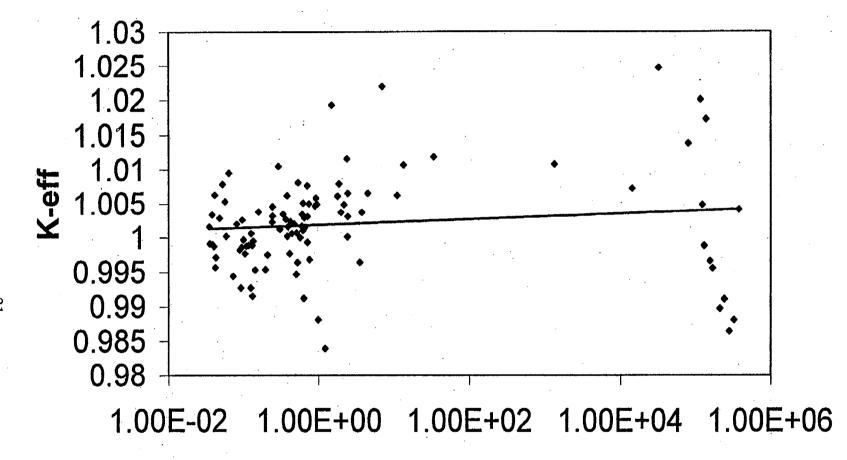


Figure 9 Trend plot for k_{eff} versus energy of average lethargy causing fission (EALF)



EALC (ev)

Figure 10 Trend plot for $k_{\it eff}$ versus energy of average lethargy causing capture (EALC)

Application of Methodology

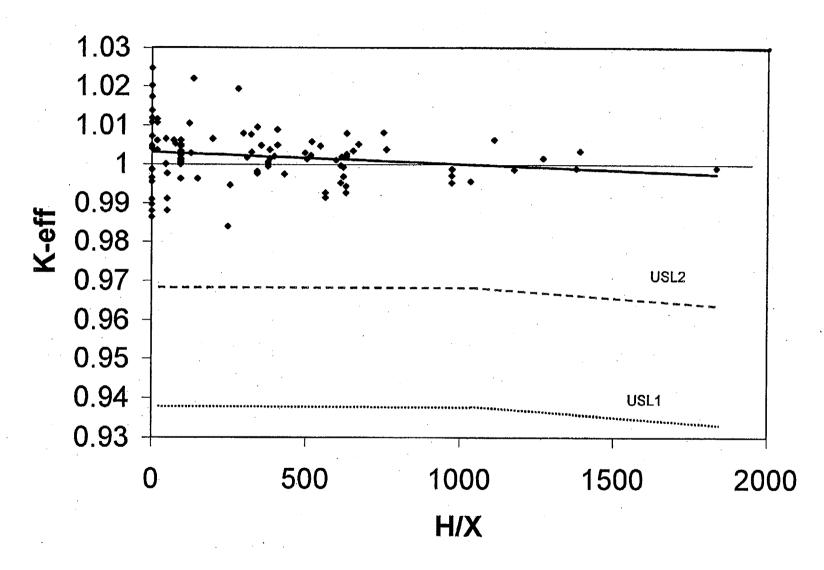


Figure 11 Trend plot for k_{eff} versus hydrogen-to-²³⁵U ratio (H/X)

Application of Methodology

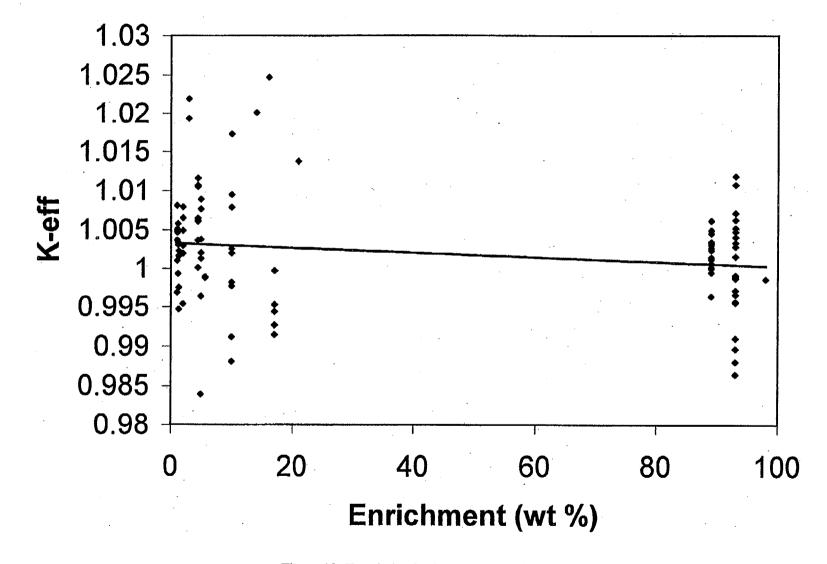


Figure 12 Trend plot for k_{eff} versus enrichment

3.4 Trending Analysis Using D Values

This section will discuss trending analyses using the same set of 102 benchmarks as the traditional analyses shown above; however, the trending parameters are now the D coefficients, described in Volume 1. Even though it is possible to perform the trending on each of the D coefficients independently, it was decided to trend k_{eff} versus the sum of these coefficients (i.e., $D_{sum} = D_c + D_n + D_s$). This combination reduces the number of trends plots to be examined.

The trend plot of $k_{\rm eff}$ versus $D_{\rm sum}$ is given in Figure 13 for the U(11)O₂ H/X=3 system.* These plots are analyzed in quite a different method from the traditional approach. Since the D coefficients reflect the deviations of the sensitivities of benchmark systems from the respective sensitivities of the application system, D equal to zero is our goal, i.e., the system does not deviate from itself. A $D_{\rm sum}$ value of zero corresponds to the U(11)O₂ H/X = 3 system for Figure 13. The trend line must therefore be extrapolated to zero in order to estimate the Δk bias. Therefore, the slope of the trend line is not nearly as important as where it crosses zero and how many systems are in the region of $D_{\rm sum}$ less than 1.2 (i.e., 3 times 0.4 as postulated earlier). Extrapolation to D = 0 should not be misinterpreted as an extrapolation to systems that are not bracketed between similar systems. From this plot it is clear that perhaps only one other system could be considered similar to the U(11)O₂ H/X = 3 system (i.e., $D_{\rm sum}$ less than 1.2). Hence, the predicted bias will have somewhat large uncertainties associated with it. The trend plot for the U(11)O₂ H/X = 40 system is shown in Figure 14. Here the coverage near a $D_{\rm sum}$ value of zero is much better than the value shown in Figure 13. In this case, there are at least 8 systems with $D_{\rm sum}$ values of 1.2 or less. The trend plot for the last U(11)O₂ system (i.e., H/X = 500) is shown in Figure 15. Here the conclusions are very similar to those in Figure 14. A large number of systems are within a $D_{\rm sum}$ value of 1.2, with a resulting good prediction of the Δk bias for this system.

These trending analysis results are generated using the same software that was used in the traditional trending approach in Section 3.3. Therefore, the same type of estimates for the Δk bias and its uncertainty can be obtained from these analyses. These bias predictions are given in Table 10. Note that the uncertainties in the predicted bias shown in Table 10 all have about the same value. This is due to the use of all benchmarks in the trend for each of these four application systems. An optional choice of using only similar experiments in the trending would have produced biases corresponding closely to the similar experiments, but would have very large (and perhaps invalid) uncertainties for the H/X = 0 and 3 cases due to the very limited number of applicable experiments.

| | | | | | | - Julia | | |
|--|----------------|---------------------------------|---------------------------------|---------------------------------|-----------------|---------------------------------|------------------|---------------------------------|
| | H/X = 0 System | | H/X = 0 System $H/X = 3$ System | | H/X = 40 System | | H/X = 500 System | |
| Trending Parameters | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} |
| $\overline{\mathrm{D}_{\mathrm{sum}}}$ | 0.74 | 0.71 | 1.20 | 0.70 | 0.19 | 0.72 | 0.09 | 0.71 |

Table 10 Predicted Δk bias and its standard deviation based on D_{sum} trending procedure

^aStandard deviations correspond to the "pooled standard deviation" as specified in Ref. 5.

^{*}The U(11)O₂ H/X = 0 system was analyzed, but the trend plot is not shown. See Table 10 for a summary of the complete results.

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NUREG/CR-6655, Vol. 2

Application of Methodology

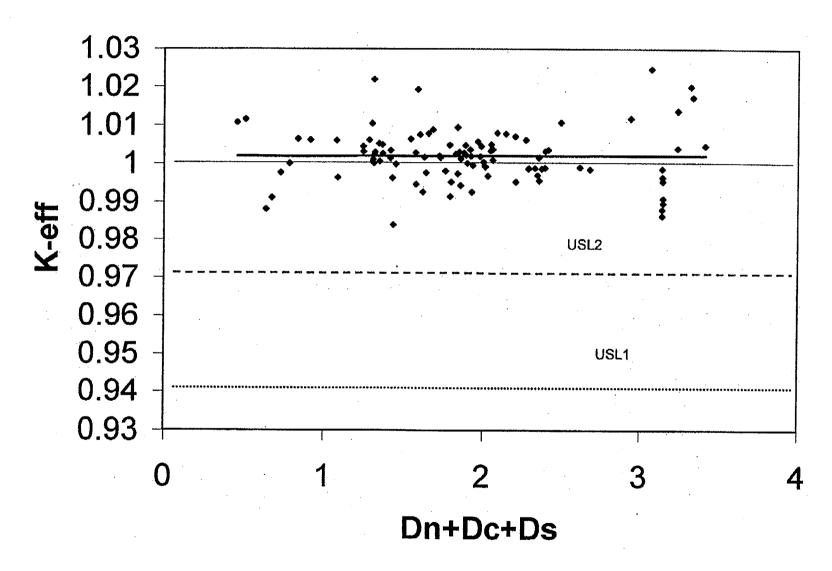


Figure 14 Trend plot for k_{eff} versus D_{sum} value for the U(11)O₂ H/X = 40 system

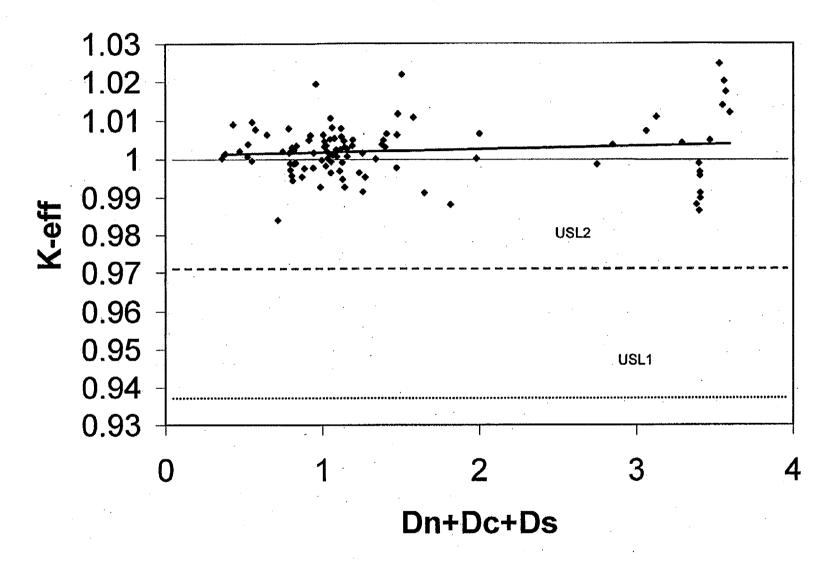


Figure 15 Trend plot for k_{eff} versus D_{sum} value for the U(11)O₂ H/X = 500 system

3.5 Trending Analysis Using c_k Values

The trending analyses using the c_k values follow very closely to the analyses using the D coefficients discussed in Section 3.4. Here the trend curves are interpreted as an extrapolation to a c_k value of unity, which corresponds to the particular application system of interest; i.e., correlation coefficient with an experiment identical to the application. The slope of the trend curve is again of secondary importance. The items of primary importance are the number of systems with a c_k value greater than 0.8 and the value of the predicted Δk bias at a c_k value of unity.

The k_{eff} trend plot for c_k of a U(11)O₂ H/X = 0 system is shown in Figure 16. This trend plot is interesting when compared with the traditional trend plot shown in Figure 9. The four data points in the upper-right-hand portion of both plots correspond to the same four systems (three ZPR and the Big-10). In Figure 9, the predicted Δk bias is about 0.4% because the overprediction of k_{eff} for these four systems is counteracted by the underprediction of the HEUMET systems, which all have very similar values of EALF. However, the trend seen for k_{eff} in Figure 16 is caused by the lack of similarity between the U(11)O₂ H/X = 0 and HEUMET systems. These HEUMET systems have a c_k value of about 0.5–0.6, indicating only minor correlations with the U(11)O₂ H/X = 0 system. This example shows the potential improvement from the use of a trending analysis with these new parameters, since trends can be observed as a function of systems that are expressly determined to be similar. It is clear from the preceding analyses that sometimes the traditional parameters indicate that systems should be similar, but are not.

The trend plots for the remaining U(11)O₂ systems with H/X values of 3, 40, and 500 are given in Figures 17 to 19. For the system with an H/X value of 3, the predicted biases are higher than those predicted by the standard techniques. The specific reasons for these differences were not explored in depth as with the H/X of 0 cases, but are believed to be caused by the separation of effects that tended to cancel each other. In any case it should be noted that all points with $c_k \ge 0.8$, i.e., similar systems, have k_{eff} values that are greater or equal to one. The Δk bias predicted for the H/X = 40 and H/X = 500 systems are in line with those of the standard techniques since there are a large number of experiments considered to be similar, and no cancellation of effects is seen.

These trending analysis results are generated using the same software that was used previously in Sections 3.3 and 3.4. Estimates of the Δk bias and its uncertainty from this trending approach are given in Table 11. Again, note the similarities of the uncertainties in the bias predictions. This is due to the use of all benchmarks in the trending, regardless of the magnitude of their c_k values.

| | H/X = 0 System | | = 0 System H/X = 3 System | | H/X = 40 System | | H/X = 500 System | |
|---------------------|----------------|---------------------------------|---------------------------|---------------------------------|-----------------|---------------------------------|------------------|---------------------------------|
| Trending Parameters | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} |
| c_k | 1.34 | 0.66 | 1.15 | 0.66 | 0.26 | 0.72 | 0.39 | 0.79 |

Table 11 Predicted Δk bias and its standard deviation based on c_k trending procedure

^aStandard deviations correspond to the "pooled standard deviation" as specified in Ref. 5.

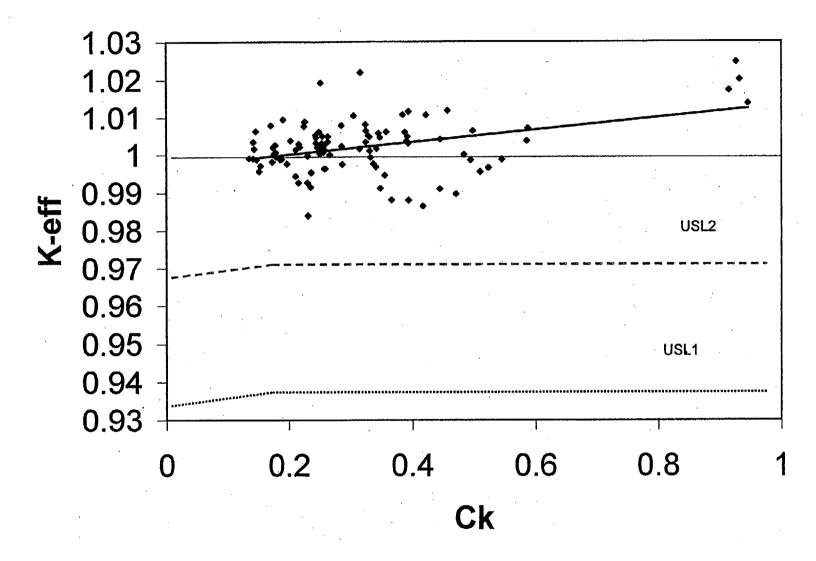


Figure 16 Trend plot for k_{eff} versus c_k value for the U(11)O₂ H/X = 0 system

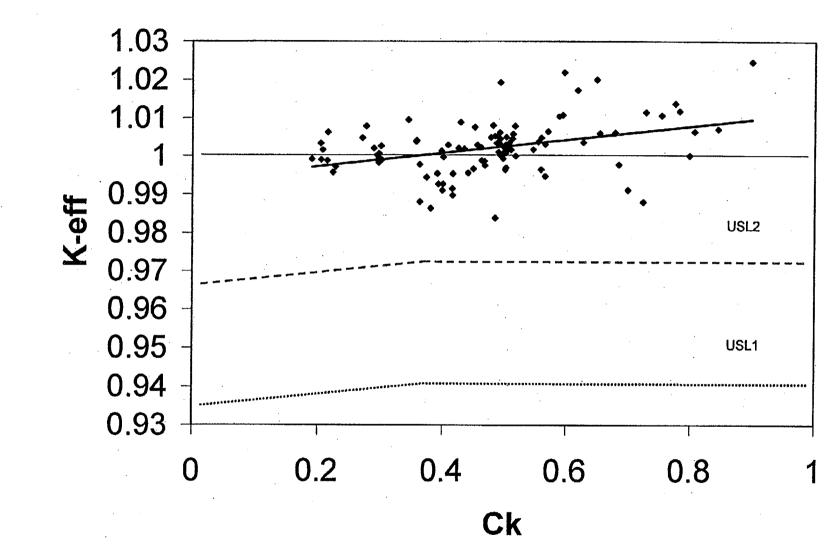


Figure 17 Trend plot for k_{eff} versus c_k value for the U(11)O₂ H/X = 3 system

Application of Methodology

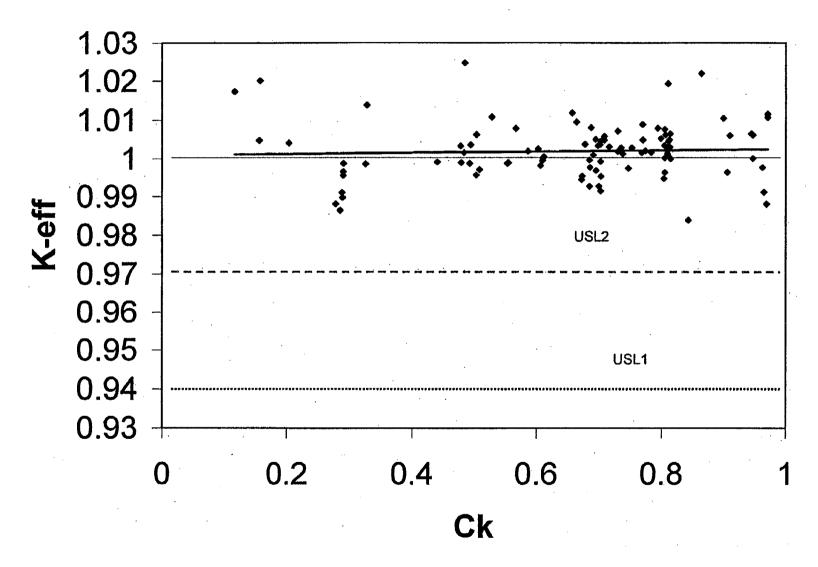


Figure 18 Trend plot for k_{eff} versus c_k value for the U(11)O₂ H/X = 40 system

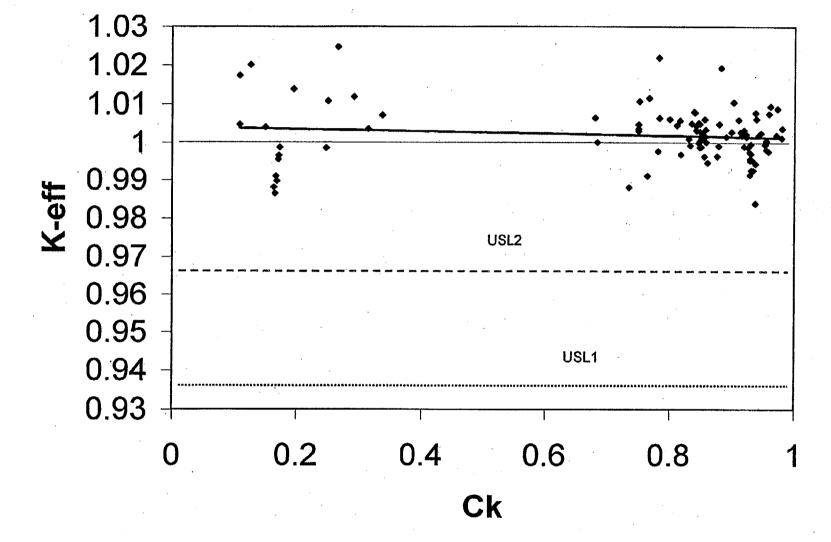


Figure 19 Trend plot for k_{eff} versus c_k value for the U(11)O₂ H/X = 500 system

3.6 GLLSM Analysis of 11-wt % Systems

The GLLSM technique is a procedure completely different from the preceding trending techniques, although it does have the same intended endpoint (i.e., the determination of the predicted Δk bias and associated uncertainty for the set of U(11)O₂ systems) and GLLSM uses the same sensitivities and correlations as the other techniques. As was shown in Volume 1 of this report, the GLLSM procedure is capable of indicating not only the predicted Δk bias, but also demonstrates its convergence to a single value as the number of experiments is increased. The (e-a)/c quantity used in Figure 20 is the same output value discussed in Volume 1; however, it has a different meaning with regard to an application area as used here. The difference in meaning is due to the lack of an experimental value, e, for an unmeasured application. By convention the calculated and measured values are set to 1.0 for the application, thus (e-a)/c is directly related to $k_a(\alpha) - m_a'$ shown in Eq. (20) in Volume 1, which is the prediction of Δk bias. Shown in Figure 20 is the convergence of the predicted Δk bias as a function of experiment groups. The experimental groups are defined based on similar system characteristics as follows:

| Group No. | Experiment descriptions |
|-----------|---|
| 1 | Godiva, HISS, HEUMET |
| 2 | ORNL spheres |
| 3 | UH3, ORNL L7-L11 |
| 4 | U(5) ₃ O ₈ , Big-10, ZPR, |
| . 5 | SHEBA, Rocky Flats |
| 6 | U(2)F ₄ , U(3)F ₄ |
| 7 | BAPL, PCTR, TRX |
| 8 | Russian experiments |

The moderated systems (H/X > 20) appear to converge quicker than the fast systems. In this example, the fast systems require four of the groups to be included for convergence, while the moderated systems only required three groups. This difference is likely due to the larger number of moderated systems, since the procedure does not favor one group of systems over the other. The predicted Δk bias and uncertainties for all 14 of these systems are shown in Table 12.

Application of Methodology

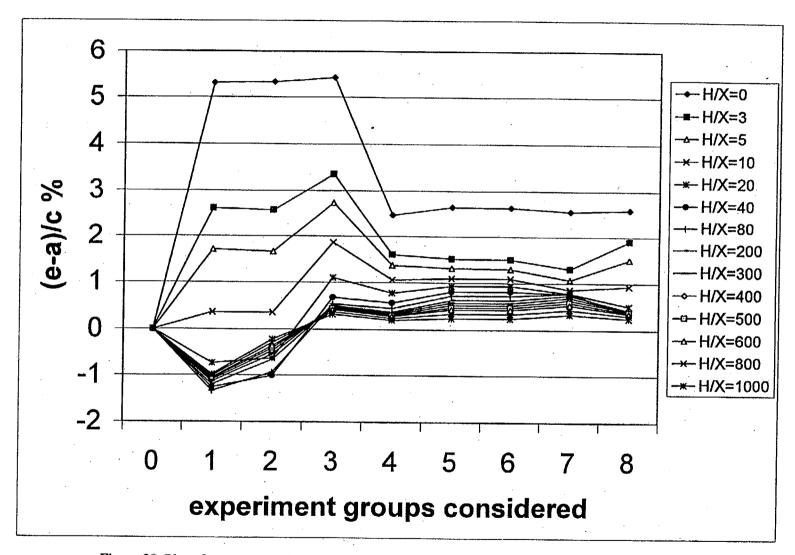


Figure 20 Plot of convergence in predicted k_{eff} bias, [(e-a)/c], for U(11)O₂ systems as a function of number of experimental groups considered

| | | | | | | * | | |
|--------|---------------------|---------------|---------------------------------|--------|------------------|-------------|---------------------------------|--|
| H/X | = 0 System | H/X = | 3 System | H/X = | = 5 System | H/X | = 10 System | |
| % bias | % std. dev. | % bias | % std. dev. | % bias | % std. dev. | % bias | % std. dev. | |
| | of biased k_{eff} | | of biased k_{eff} | | of biased k | aff | of biased k_{eff} | |
| 2.59 | 0.46 | 1.90 | 0.39 | 1.50 | 0.35 | 0.93 | 0.35 | |
| | | | | | | | | |
| H/X = | = 20 System | H/X = | 40 System. | H/X = | 80 System | H/X = | = 200 System | |
| % bias | % std. dev. | % bias | % std. dev. | % bias | % std. dev. | . % bias | % std. dev. | |
| | of biased k_{eff} | | of biased k_{eff} | | of biased k | eff | of biased $k_{e\!f\!f}$ | |
| 0.51 | 0.39 | 0.35 | 0.43 | 0.36 | 0.46 | 0.42 | 0.46 | |
| | | | | | | | | |
| H/X = | 300 System | H/X = | H/X = 400 System | | H/X = 500 System | | = 600 System | |
| % bias | % std. dev. | % bias | % std. dev. | % bias | % std. dev | . % bias | % std. dev. | |
| | of biased k_{eff} | , | of biased k_{eff} | | of biased k | eff | of biased k_{eff} | |
| 0.41 | 0.42 | 0.40 | 0.39 | 0.37 | 0.35 | 0.35 | 0.33 | |
| | | | | | | | | |
| | H/X = 3 | 800 System | | | H/X = | 1000 System | | |
| % t | oias | % std. dev. o | % std. dev. of biased k_{eff} | | % bias | | % std. dev. of biased k_{eff} | |
| 0.30 | | 0.2 | 29 | 0.24 | | (| 0.28 | |

Table 12 Predicted Δk bias and its standard deviation^a based on GLLSM procedure

3.7 Summary of Trending Analyses

In the preceding sections, results from a number of approaches to criticality safety data validation were presented. Quite interestingly, they give very different answers for the set of application problems chosen for study. The predicted bias are in some cases up to a factor of 15 different. The primary reason for these differences seems to be the inclusion of systems that may "look" very similar from the standpoint of certain parameters, but vary with respect to other parameters. In particular, according to both the H/X and EALF parameters, the HEUMET and ZPR/Big-10 problems are similar. However, with respect to the sensitivities and uncertainties, they appear to be quite different. Cancellation of effects due to systems that "appear" to be similar causes the traditional trending approaches to underpredict the actual bias for low-moderation systems with intermediate enrichments. This situation is evident in Table 13, where each of the previously reported results are presented in summary form. Note also that the differences between the c_k , D, and GLLSM results for low H/X systems would have been minimized if only the similar experiments had been included in the trending studies. The validity for inclusion and the best approach options for inclusion of dissimilar ($c_k < 0.8$ or D > 1.2) experiments in the trending analyses will be studied in future work.

[&]quot;Standard deviations correspond to the value output from the GLLSM process. See Appendix B of Volume 1 for details.

Table 13 Comparison of predicted Δk bias and the standard deviation of the estimated k_{eff} for various procedures

| | H/X = | = 0 System | H/X = 3 System | | H/X = | : 40 System | H/X = 500 System | |
|------------------------|--------|---------------------------------|----------------|---------------------------------|--------|---------------------------------|------------------|---------------------------------|
| Trending Procedures | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} | % bias | % std. dev. of biased k_{eff} |
| EALF | 0.17 | 0.70 | 0.26 | 0.70 | 0.26 | 0.70 | 0.26 | 0.70 |
| H/X | 0.32 | 0.71 | 0.32 | 0.71 | 0.31 | 0.71 | 0.17 | 0.71 |
| D_{sum} | 0.74 | 0.71 | 1.20 | 0.70 | 0.19 | 0.72 | 0.09 | 0.79 |
| C_k | 1.34 | 0.66 | 1.15 | 0.66 | 0.26 | 0.72 | 0.39 | 0.79 |
| GLLSM | 2.59 | 0.46 | 1.90 | 0.39 | 0.35 | 0.43 | 0.37 | 0.35 |

[&]quot;Standard deviations correspond to the "pooled standard deviation" as specified in Ref. 5.

The predicted bias from these applications are all positive. Since the USLSTATS program⁵ sets a positive bias to zero for the purposes of determining USL1 and USL2, the overprediction found in these applications does not present a safety concern. However, a similar situation can be easily postulated where a predicted positive bias is actually a negative bias. With the inclusion of strict confidence levels, along with an additional margin of subcriticality, the cumulative effect of these factors should still be conservative. However, prudent application of trending procedures is very important in criticality safety validation exercises.

It can be seen in Table 14 that the predicted USL1 (with an assumed 5% administrative margin of subcriticality for safety) and USL2 values are nearly identical for all the procedures, with the exception of the GLLSM, which predicts USL values some 0.5-1.5% higher due to the lower uncertainty estimates. Caution is advised with the direct use of the USL values from the GLLSM since these results are estimated based on a comparison of the standard deviations and USL values from the c_k trending analysis. The development of the statistical basis for the USL values using the GLLSM approach is currently under way in another related project.

Table 14 Comparison of predicted USL1 and USL2 values for the various procedures

| Trending | H/X = 0 System | | H/X = 3 System | | H/X = 40 System | | H/X = 500 System | |
|--------------------|----------------|-------|----------------|-------|-----------------|-------|------------------|-------|
| Procedures | USL1 | USL2 | USL1 | USL2 | USL1 | USL2 | USL1 | USL2 |
| EALF | 0.938 | 0.970 | 0.938 | 0.970 | 0.938 | 0.970 | 0.938 | 0.970 |
| H/X | 0.937 | 0.969 | 0.937 | 0.969 | 0.937 | 0.969 | 0.937 | 0.969 |
| \mathbf{D}_{sum} | 0.937 | 0.968 | 0.937 | 0.967 | 0.938 | 0.970 | 0.938 | 0.971 |
| c_k | 0.938 | 0.971 | 0.938 | 0.971 | 0.937 | 0.969 | 0.936 | 0.966 |
| GLLSM | 0.942 | 0.979 | 0.943 | 0.982 | 0.942 | 0.982 | 0.944 | 0.985 |

4 GUIDANCE FOR DETERMINATION OF APPLICABILITY AND ADVANCED TRENDING PROCEDURES

The proposed use of multiple criticality safety data validation techniques is not to make the task of validation more difficult, neither is it to make the task less difficult. The goal of this work is to allow for criticality safety validation under conditions not possible previously, primarily the extrapolation beyond the range of the experimental benchmarks. Techniques developed in this work provide a more rigorous and formal definition of what applicability really connotes.

4.1 Recommendations on the Use of Trending Procedures

As stated in the Volume 1 summary of this document, 6 this work has recommended that two systems with a c_k value of 0.8 or higher can be considered similar enough that one can serve as a benchmark in the validation of the other. This conclusion does not imply that a formal S/U analysis needs to be redone for all previous criticality validation exercises. Clearly, a series of systems with the nearly identical characteristics to the application of interest (i.e., same enrichments, same moderators, a very tightly spaced sequence of moderator/fuel ratio values, similar geometries, etc.,) should be acceptable for use as a validation set. Indeed, in Volume 1 such a series of experiments with constant enrichment was shown to have c_k values greater than 0.9 for H/X values within a factor of 2 of each other.

Thus, the recommendation is that for benchmark data sets that contain more than 20 benchmarks with a c_k value of 0.8 or higher to the corresponding application area, a traditional criticality safety data validation via the standard k_{eff} trending analysis method is acceptable. This recommendation doubles the Volume 1 predicted numbers of benchmarks needed for a reasonable estimate of the k_{eff} bias or the resulting upper subcritical limit. It is further recommended that in using the traditional k_{eff} trending approach that systems with very different characteristics such as high-versus-low enrichment, poisoned-versus-nonpoisoned, wet-versus-dry, etc., not be included together unless that specific effect is being analyzed. For example, do not include three or four random systems with boron poisons in an analysis of nonpoisoned systems. As was seen in the comparison of the traditional-versus-advanced validation procedures in Table 12, the inclusion of widely varying benchmarks can cause cancellation of trends, resulting in an under prediction of the k_{eff} bias. Similar situations could also result in an over prediction of the k_{eff} bias.

The advanced trending analyses, using the D and c_k coefficients as trending parameters, are recommended under conditions that are similar but complimentary to the above recommendations for the traditional trending analyses. The conditions on the advanced trending analyses parameters are as stated in Volume 1 for a meaningful estimate of the k_{eff} bias (i.e., 5–10 experiments with a c_k value of 0.9 or higher or 10–20 experiments with a c_k value of 0.8 or higher). These recommended parameter ranges are only marginally different from those above; however, the motivation for use of these advanced trending analyses is that the inclusion of widely varying system conditions like H/X, enrichment, poisons, etc., are actually encouraged. These advanced techniques are very adept at differentiating between dissimilar systems, while allowing for a wide range of system types to be validated using the single validation data set. To date, the practice has been to include all systems in the benchmark set in the trending analyses; however, future work is planned to explore using only those experiments that are determined to be similar. Applications for which the advanced validation techniques should be applicable are extrapolation scenarios, which are necessary when the desired experiments meeting the strict requirements of the standard techniques are not available. This set of recommended application areas is very well represented by the examples shown in Section 3 for the greater-than-5-wt % enrichment cases.

The final data validation technique presented in this work is the GLLSM methodology. This technique has been extremely useful in the prediction of values of c_k and D that are necessary to define an area of applicability, as well as the determination of the number of experiments needed to obtain a valid estimate of the k_{eff} bias and associated uncertainty. Additionally, this technique can be a useful data validation procedure under practically any set of conditions. However, since the technique is currently much more complex to implement than the other procedures, it is recommended that it only be used in situations where the other approaches are not expected to proceed satisfactorily. The GLLSM technique corresponds to a general validation, where the values of c_k and D fall outside the ranges specified above, or when responses other than k_{eff} are needed for inclusion into the validation. Using the full GLLSM techniques for criticality safety validation can allow for noncritical experiments to be included into the validation procedures. Further development is required before the full GLLSM procedure can be used in these applications, since the "similarity" criteria for the c_k and D parameters may be different when using noncritical systems. A means for ensuring that all important material/reactions and energy regions of a given application problem are tested by the validation procedure must be developed. This procedure would then allow for validation to occur even with a very diverse set of benchmark data (e.g., noncritical measurements).

4.2 Trending Procedures for c_k and D Parameters

The techniques for doing trending analyses with the c_k and D parameters are quite different from those of traditional trending analyses. The traditional analyses allow for multiple systems to be validated from a single trending curve, while the trends with c_k and D parameters only allow a single system per trend plot to be validated. This limitation results from the fact that the trend parameters are only valid between pairs of systems. Thus the procedure for doing a trend analysis, with either the c_k or D parameters, is to first generate sensitivity coefficients for all the benchmark problems and all the application scenarios. The application scenarios should cover the range of materials, moderation, enrichment, impurities, etc., that are expected to be used in the criticality safety analyses. If a broad range is needed, it is recommended that the range be broken up into discrete values (like the H/X values in the U(11)O₂ example above). A simple postprocessing code should then be able to generate either the c_k or D parameters for all systems. These parameters can then be trended using standard data-fitting codes with the resulting bias and uncertainties of fit and future predictions being extrapolated to either the zero parameter value for D or the unity parameter value for c_k . This trending must be done for each of the application systems that span the range of scenarios for the criticality safety analyses.

The above techniques specify that the sensitivities for all of the benchmark and application systems be quantified. Currently, only 1-D and 2-D geometry capabilities exist for the generation of this sensitivity information; therefore, it is necessary to convert the critical systems into either of these two geometrical systems. Standard procedures are available that make this conversion process easier (cell-weighting, equivalent buckling, etc.), but the conversion can be a time-consuming process that requires some degree of expertise. The development of 3-D sensitivity tools should help alleviate this limitation.

4.3 Guidance on Estimating Experimental Uncertainties

The estimate of an uncertainty matrix can be quite difficult to construct, depending on the availability of a complete description of the experiment. Experimentalists routinely give estimates of the uncertainty in a single experiment, but do not always give enough information to estimate the common uncertainties between a series of critical measurements. If the number of experiments is relatively large, the detailed uncertainty information is typically not

needed. Under these circumstances, it is often desirable to assume that the uncertainties in the respective experiments are uncorrelated and allow for the experimental uncertainties to vary between maximum and minimum quantities. This variation in uncertainties allows a test of the variation in the bias results to be determined. The remainder of this section describes how a full matrix of uncertainties can be generated if sufficient experimental detail is available.

The report describing the $U(2)F_4$ and $U(3)F_4$ experiments¹¹ estimates the uncertainties in the individual material number densities. The report gives standard deviations in the UF₄ weight percent of 0.7%, paraffin weight percent of 1.2%, enrichment of 1.0% (0.7% for the 3 wt % cases), and density correction of 0.3%. These uncertainty contributions can be used to define a covariance matrix (×10⁴) corresponding to the individual isotopes for the 2 wt % cases as follows:

| Material | ²³⁵ U | ²³⁸ U | Н | С | F |
|------------------|-------------------------|------------------|------------------|-----------------|-----------------|
| ²³⁵ U | $0.7^2 + 1.0^2 + 0.3^2$ | $0.7^2 + 0.3^2$ | 0.3 ² | 0.3^{2} | $0.7^2 + 0.3^2$ |
| ^{238}U | $0.7^2 + 0.3^2$ | $0.7^2 + 0.3^2$ | 0.3^{2} | 0.3^{2} | $0.7^2 + 0.3^2$ |
| H | 0.3^{2} | 0.3^{2} | $1.2^2 + 0.3^2$ | $1.2^2 + 0.3^2$ | 0.3^{2} |
| С | 0.3^{2} | 0.3^{2} | $1.2^2 + 0.3^2$ | $1.2^2 + 0.3^2$ | 0.3^{2} |
| F | $0.7^2 + 0.3^2$ | $0.7^2 + 0.3^2$ | 0.3 ² | 0.32 | $0.7^2 + 0.3^2$ |

These covariances can then be combined with the respective sensitivities of k_{eff} to the number densities for each isotope to obtain the uncertainty in k_{eff} due to these effects. The resulting k_{eff} standard deviations for each of the eight experiments are 0.49, 0.45, 0.46, 0.44, 0.42, 0.45, 0.53, 0.52%, respectively. The remaining six experiments for U(5)₃O₈ are constructed using the same methods.

4.4 Future Developments

Volume 1 of this report presented the development of S/U and GLLSM techniques for application to criticality safety data validation studies. This volume has applied those methods to the analysis of uranium systems with enrichments greater than 5 wt %. Specifically, a family of $U(11)O_2$ systems with H/X values of 0–1000 have been analyzed. The guidelines given in the previous sections explain when usage of traditional procedures is recommended as well as the situations when the use of these newly developed procedures is recommended. The procedures for using the 1-D sensitivity tool, SEN1, and for trending bias with the resulting D and c_k parameters are established as a result of this development work, however, additional experience with other application areas is required with these new procedures before complete guidance on their use is possible. Work is planned to address the trending with D and c_k values which are outside the proposed similarity criterion of this report. Comparison of trending procedures which exclude experiments which are not similar, include them as shown in this work, and weight them based on their c_k and D values are exercises planned for the future. This experience should allow for development of an upper subcritical limit (USL) procedure that completely takes advantage of these new validation methods. For example, the possibility of a lower administrative margin due to the inclusion of cross section uncertainty effects could be explored.

There are a number of development areas that need further study regarding the use of these techniques. The first area is the need for flexible tools for applying the sensitivity methods to 2-D and 3-D problems. A 2-D capability is currently available,²¹ but further work on implementation of a more flexible package is planned. The use of existing inputs for a 3-D Monte Carlo code is considered a must for the routine application of these procedures. Currently work is underway on this task,²⁴ but further work is needed before this tool will be generally available for use.

The sensitivity methodology applied in this work obtains the sensitivity of $k_{\rm eff}$ to the problem-specific cross sections used in the analysis. Future work is also needed to explore extension of the sensitivity capability to obtain the sensitivity of $k_{\rm eff}$ to the code and data parameters that impact the problem-specific cross section processing. Reference 20 indicates via direct re-calculations how omission of these effects can cause errors in the sensitivity coefficients that are obtained.

The procedures that govern the use of the GLLSM techniques require some additional development before they can be routinely applied in a criticality safety data validation process. These additional developments involve a concept called "completeness," the determination of the rigor involved for data validation versus data evaluation (i.e., self-shielding effects, inclusion of resonance parameter uncertainties, etc), and the addition of non-critical experiment and/or non- k_{eff} response methods.

The completeness concept is needed to ensure that in the general application of GLLSM techniques, where there are no limits for the values of c_k and D, the benchmark data set is complete in the sense that all important reactions and energies are validated. There is a recognized need to investigate and address, as appropriate, the influence and magnitude of shielded and unshielded effects of resonance reactions on the sensitivity and uncertainty analyses relating the GLLSM procedure. The resonance effects include the option of incorporating uncertainties due to self-shielding into the analysis, and also the generation of and processing of resonance parameter uncertainties. The use of non-critical and/or non- k_{eff} responses in the GLLSM analysis is currently available, however, it is not in the SEN1 package that was developed for this project. Either the capability could be added to the SEN1 package or currently available programs that have more complex input requirements could be used. The full use of these capabilities in a 3-D sensitivity generation procedure would simplify the criticality safety validation efforts. Additional measures are also expected to be needed in order for the full GLLSM techniques to be readily used by criticality safety practitioners.

The trending procedures with c_k and the GLLSM applications require the availability of cross section covariance files for processing to a desired group structure. There are currently 27 materials/isotopes with covariance data in ENDF/B-VI. However, there are still many important nuclides without covariance information. This information needs to be generated for many isotopes that are of importance to the various criticality application areas.

While the methods and procedures generated as a result of this project are felt to be quite general, they have currently only been applied to uranium systems. Application to other fissile material (e.g., plutonium or mixed oxide) systems is desired to ensure the generality of these procedures. For systems containing primarily uranium fissile materials, the D_{sum} and c_k parameters appear to work equally well. Application to other fissile materials should allow for confirmation of these trends.

5 SUMMARY

This report presents the application of sensitivity and uncertainty (S/U) analysis methodologies developed in Volume 1 to the code/data validation tasks of a criticality safety computational study. Sensitivity and uncertainty analysis methods were first developed for application to fast reactor studies in the 1970s. This work has revitalized and updated the existing S/U computational capabilities such that they can be used as prototypic modules of the SCALE code system, which contains criticality analysis tools currently in use by criticality safety practitioners. After complete development, simplified tools are expected to be released for general use.

The methods for application of S/U and generalized linear-least-squares methodology (GLLSM) tools to the criticality safety validation procedures were described in Volume 1 of this report. Volume 2 of this report presented the application of these procedures to the validation of criticality safety analyses, supporting uranium operations for use in commercial power reactors where the enrichments are greater than 5 wt %.

As a part of this study, a benchmark database of 102 experiments was developed which included not only calculational models and measured-versus-calculational $k_{\rm eff}$ results but also sensitivity and uncertainty results generated by the S/U methodology described in Volume 1. The same tools were then applied to a set of 14 application scenarios corresponding to U(11)O₂ systems with H/X values ranging from 0 to 1000. A range of applicability determination was performed using the D and c_k coefficients generated by the S/U analysis. The analysis indicated that the benchmark dataset had good coverage of the 11-wt % systems with H/X values greater than 40, while the systems with H/X less than 40 had only marginal-to-inadequate coverage. The sparsity of low-H/X systems is a concern. Additional experiments in this area would be useful.

Additionally, analyses were performed to predict biases for systems with enrichments greater than 5 wt %. Specifically, the traditional k_{eff} trending analyses were compared with newly developed k_{eff} trending procedures, utilizing the D and c_k coefficients described in Volume 1. Application of these newly developed procedures was made to the family of applications involving U(11)O₂ systems with H/X values ranging from 0 to 1000.

The comparisons among the various trending techniques were quite interesting in that they give very different answers, depending on the particular system being analyzed. The predicted biases for various systems were in some cases up to a factor of 5 different between the various trending techniques. The primary reason for these differences was that various systems "look" very similar from the standpoint of certain parameters, but are very different with respect to other parameters. In particular, the H/X and EALF parameters predicted similarity between dry systems with high enrichments (93 wt %) and dry systems with intermediate enrichments (10 wt %), while the c_k and D parameters indicated that these systems were quite different. The net effect of trending with H/X and EALF was a cancellation of effects (the HEU systems were underpredicted by about 1%, the IEU systems were overpredicted by about 2%), which produced a predicted bias of about 0.5% overprediction for a dry U(11)O₂ system. The trending with c_k and D parameters produced an estimated bias of between 1 and 2% overprediction since only the IEU systems were predicted to be similar to the dry U(11)O₂ system. Although the predicted biases from these applications are all positive, the differences in magnitude are a concern, since the prudent application of trending procedures is very important in criticality validation exercises.

Summary.

Section 5

As a result of these observations, the following guidance was developed for usage of the various trending techniques.

- Traditional trending procedures are acceptable for bias estimation if about 10–20 benchmarks with a ck value of 0.8 or higher are included in the benchmark database. It is further recommended that in using the traditional keff trending procedures that systems with very different physical characteristics, such as high-versus-low enrichment, poisoned-versus-nonpoisoned, wet-versus-dry, etc., not to be included together unless that specific effect is being analyzed.
- The advanced trending analyses with the D and c_k coefficients are recommended if fewer experiments than those recommended above are available. Also the primary motivation behind their use is that widely varying systems can be included in the analysis, hence making the effective extrapolation of the existing data more meaningful. The procedure will allow for automatic collection of experiments that have similar D and c_k coefficients such that the predicted k_{eff} bias should be valid.
- The GLLSM procedure is recommended if existing experiments with c_k values of 0.8 or higher with respect to the area of application are not available. These experiments, possibly having c_k values 0.2–0.6, are not similar to the application; but they are still relevant. A suite of such experiments can be expected to perform quite well using a GLLSM technique if they are suitably chosen to highlight different aspects of the given application system. Under these conditions, the use of GLLSM for validation is possible, although it may still be desirable to initiate an experimental program to obtain additional data. This procedure also allows for the inclusion of noncritical experiments and non- k_{eff} responses in the validation exercises. Further development is required before these GLLSM techniques can be fully functional in a routine criticality safety validation study.

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| capabilities such that they can be used as prototypic modules of the S | SCALE code system, which contains criticality analysis tools |
| currently in use by criticality safety practitioners. After complete devel | lopment, simplified tools are expected to be released for |
| general use. | |
| The methods for application of S/U and generalized linear-least square | res methodology (GLLSM) tools to the criticality safety |
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